




























Euclid preparation

LXXXVII. Non-Gaussianity of two-point statistics likelihood: Precise analysis of the matter power spectrum distribution

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ABSTRACT

We investigate the non-Gaussian features in the distribution of the matter power spectrum multipoles. Using the COVMOS method, we generated 100 000 mock realisations of dark matter density fields in both real and redshift space across multiple redshifts and cosmological models. We derived an analytical framework linking the non-Gaussianity of the power spectrum distribution to higher-order statistics of the density field, including the trispectrum and pentaspectrum. We explored the effect of redshift-space distortions, the geometry of the survey, the Fourier binning, the integral constraint, and the shot noise on the skewness of the distribution of the power spectrum measurements. Our results demonstrate that the likelihood of the estimated matter power spectrum significantly deviates from a Gaussian assumption on non-linear scales, particularly at low redshift. This departure is primarily driven by the pentaspectrum contribution, which dominates over the trispectrum at intermediate scales. We also examined the impact of the finiteness of the survey geometry in the context of the *Euclid* mission, and we find that both the shape of the survey and the integral constraint amplify the skewness.

Key words. cosmological parameters – cosmology: observations – cosmology: theory – large-scale structure of Universe

1. Introduction

In recent years, observational cosmology has experienced significant advancements with the emergence of next-generation galaxy surveys designed to probe larger cosmic volumes and smaller scales than ever before. Examples of such surveys include the *Euclid* mission (Euclid Collaboration: Mellier et al. 2025), the Vera C. Rubin Observatory Legacy Survey of Space and Time (LSST, Ivezić et al. 2019), and the Dark Energy Spectroscopic Instrument (DESI, DESI Collaboration 2016). These cutting-edge surveys will provide unprecedented insights into the structure and evolution of the Universe, making it increasingly important to develop more accurate and robust methodologies for data analysis. As we push the boundaries of observational capabilities, it is crucial that our methods of parameter inference can ensure that we can effectively extract valuable information from the wealth of data that will be gathered by these surveys.

Cosmological information is generally extracted from the analysis of two-point statistics of the fields under study. In real space, this is done with the two-point correlation function, and in Fourier space, it is done with the power spectrum. The latter is the case for state-of-the-art analyses of the temperature and polarisation of the cosmic microwave background (CMB; Planck Collaboration VI 2020) or the large-scale structure (LSS) of the Universe, using galaxy clustering obtained through spectroscopic data (Gil-Marín et al. 2020; Bautista et al. 2021; DESI Collaboration 2025) or a combination of weak lensing and clustering obtained through photometric data (Abbott et al. 2022; Asgari et al. 2021; More et al. 2023).

To extract this cosmological information, we need to assume, a priori, what the distribution of our data is in order to model the likelihood function. A simple assumption that is commonly made for two-point statistics is that they follow a multivariate Gaussian distribution. However, we know that it is not the case, even for Gaussian fields such as the CMB. Indeed, the two-point statistics of a Gaussian field follows a χ^2 distribution with a number of degrees of freedom equal to the number of independent

modes at the considered scale. On large scales this number is low such that the χ^2 distribution is far from being a Gaussian (Hamimeche & Lewis 2008).

Moreover, at intermediate scales (above 30 Mpc/h), such as those probed by the *Euclid* mission (Euclid Collaboration: Mellier et al. 2025), the non-linear clustering of matter generates non-vanishing N -point statistics (with $N > 2$) of the density field. A well-known consequence of the non-Gaussianity of the density field from the non-linear clustering of matter on the two-point statistics is the non-Gaussian contribution to the covariance matrix from a non-vanishing four-point correlation function (Scoccimarro et al. 1999). This higher-order contribution to the covariance is still present even though the likelihood is approximated to be Gaussian. In addition, these higher-order correlation functions also induce higher-order moments in the distribution of the two-point statistics, thus causing departure from the Gaussian assumption also at small scales. Indeed, the skewness and kurtosis of the two-point statistics distribution have been shown to be non-vanishing on small scales and especially at low redshift (Takahashi et al. 2009; Blot et al. 2015; Lin et al. 2020; Upham et al. 2021). Hall & Taylor (2022) also studied the potential influence of higher-order N -point statistics on the non-Gaussian distribution of two-point statistics on large scales in the context of a weak lensing likelihood analysis, but they found it to be negligible.

Given the precision with which we aim to measure the cosmological parameters with *Euclid*, we need to understand and control in detail the likelihood modelling and test whether it can significantly bias our cosmological analyses. Indeed, cosmological constraints derived from likelihood analysis generally rely on the assumption that the measurements are distributed according to a Gaussian. As a result, if the Gaussian assumption is not justified, it might affect the derived cosmological parameters. This is the aim of the present paper and its companion, Gouyou Beauchamps et al. (in prep., hereafter cited as GB-2025). In this work, we focus particularly on the Fourier version of two-point statistics, the power spectrum, to develop an analytical framework to understand how the non-Gaussianity of the two-point

statistics distribution is linked to the non-Gaussianity of the field itself and what are the settings of the likelihood to which it is the most sensitive. In particular, we look at how the skewness and kurtosis of the distribution of the matter power spectrum changes with respect to (i) cosmological effects such as the cosmological model, the redshift, and the presence or absence of redshift space distortions (RSD); (ii) the settings of the power spectrum estimator, such as the presence of aliasing and the Fourier mode binning; and (iii) the effects originating from specific features of the survey, such as the sample density and the survey footprint. Our companion paper, GB-2025, specifically focuses on the impact of wrongly assuming a Gaussian likelihood on *Euclid* cosmological constraints.

To achieve all of this, we extensively rely on the COVMOS method (Baratta et al. 2020, 2023) in order to produce very large sets of approximate simulations of non-Gaussian fields, which allow us to precisely explore the distribution of the power spectrum. It has been shown in Baratta et al. (2023) that COVMOS is a reliable tool to model the covariance of two-point statistics, and we show in this paper that it is also the case for the study of their full distribution.

Since the main scope of the present paper is to understand the mechanisms possibly producing a non-Gaussian behaviour of the distribution of the estimated power spectrum, we focus on the matter field. In GB-2025, we use more realistic mock galaxy catalogues to study the skewness of the galaxy power spectrum. Finally, we note that this work does not aim to study the non-Gaussianity of the likelihood arising from the estimation of a covariance matrix with a low number of realisations (Sellentin & Heavens 2016; Percival et al. 2022).

The paper is structured as follows. In Sect. 2 we present our analytical framework to study the distribution of the power spectrum. In Sect. 3 we describe the simulation sets we constructed with COVMOS and introduce the estimators we applied on them. In Sect. 4 we study the estimated distribution of the power spectrum under various scenarios. In Sect. 5 we study the distribution of individual Fourier modes. Finally, we conclude in Sect. 6.

2. Skewness of the power spectrum estimator

In this section we derive the skewness of the probability distribution of the estimator of the power spectrum in terms of the higher-order N -points cumulant moments of the density field. We start with some definitions. The ensemble average of the random variable X following the probability density function (PDF) \mathcal{P} is $\langle X \rangle$, and $\langle X \rangle_c$ represents its connected expectation value. In practice, the random variable we study is the estimator of the multipoles of the power spectrum, $\hat{P}^{(\ell)}(k)$, defined as

$$\hat{P}^{(\ell)}(k) := k_F^3 \frac{2\ell + 1}{N_k} \sum_{i=1}^{N_k} |\delta_{k_i}|^2 \mathcal{L}_\ell(\mu_i), \quad (1)$$

where ℓ refers to the order of the considered multipole, \mathcal{L}_ℓ are the Legendre polynomials of order ℓ , k is the modulus of the considered spherical k -shell in Fourier space, N_k is the number of independent modes that it contains, μ_i is the cosine of the angle between the line of sight and the wave vector k_i and the sum is made over independent modes inside the k -shell. Note that, by independent we do not mean statistically independent, but we rather refer to the fact that since the configuration space density field $\delta(\mathbf{x})$ is real, its Fourier space counterpart δ_k satisfies $\delta_{-k} = \delta_k^*$, where the \star stands for complex conjugate. This means that there are repeated modes, that are not considered in

Eq. (1). Unless otherwise stated, we consider a periodic boundary¹ conditions of size L , thus characterised by its fundamental mode $k_F := 2\pi/L$. That is the reason why the sum in Eq. (1) is discrete. Finally, due to statistical invariance by translation (i.e. statistical homogeneity) one can define the power spectrum (P), the bispectrum (B), the trispectrum (T), and the pentaspectrum (Q) as

$$k_F^3 \langle \delta_{k_1} \delta_{k_2} \rangle_c = \delta_{k_{12}}^K P(\mathbf{k}_1), \quad (2)$$

$$k_F^3 \langle \delta_{k_1} \delta_{k_2} \delta_{k_3} \rangle_c = \delta_{k_{123}}^K B(\mathbf{k}_1, \mathbf{k}_2), \quad (3)$$

$$k_F^3 \langle \delta_{k_1} \delta_{k_2} \delta_{k_3} \delta_{k_4} \rangle_c = \delta_{k_{1234}}^K T(\mathbf{k}_1, \mathbf{k}_2, \mathbf{k}_3), \quad (4)$$

$$k_F^3 \langle \delta_{k_1} \delta_{k_2} \delta_{k_3} \delta_{k_4} \delta_{k_5} \delta_{k_6} \rangle_c = \delta_{k_{123456}}^K Q(\mathbf{k}_1, \mathbf{k}_2, \mathbf{k}_3, \mathbf{k}_4, \mathbf{k}_5), \quad (5)$$

where we use the specific notation $\delta_{k_{1..N}}^K$ for the Kronecker symbol, which is 1 if $\mathbf{k}_1 + \dots + \mathbf{k}_N = \mathbf{0}$ and 0 otherwise. To simplify the following calculations, we defined the random variables $X := \hat{P}^{(\ell)}(k)$, $X_i := |\delta_{k_i}|^2$ and the coefficients $a_i := k_F^3 \frac{2\ell+1}{N_k} \mathcal{L}_\ell(\mu_i)$. Thus, X_i depends on the wave vector k , and X depends on the wave mode k and on the chosen multipole ℓ . Based on these definitions, Eq. (1) takes the form

$$X = \sum_{i=1}^{N_k} a_i X_i, \quad (6)$$

Since the multi-point probability distribution function $\mathcal{P}(X_1, \dots, X_{N_k})$ of the modes X_i is not known, one can only focus on the cumulant moments in order to characterise how far from a Gaussian the distribution of X is. Indeed, if the variable X were following a Gaussian distribution then its cumulant moments should be $\langle X^n \rangle_c = 0$ for $n \geq 3$. Taking the ensemble average of Eq. (6) one can show that the mean of X is given by

$$\begin{aligned} \langle X \rangle &= \frac{2\ell + 1}{N_k} \sum_{i=1}^{N_k} P(\mathbf{k}_i) \mathcal{L}_\ell(\mu_i) \\ &\simeq \frac{2\ell + 1}{2} \int_{-1}^1 P(\mathbf{k}) \mathcal{L}_\ell(\mu) d\mu = P^{(\ell)}(k), \end{aligned} \quad (7)$$

which means that the estimator is unbiased in the large N_k limit. We can then show that its variance is given by

$$\langle X^2 \rangle_c = \sum_{i=1}^{N_k} a_i^2 \langle X_i^2 \rangle_c + 2 \sum_{j>i} a_i a_j \langle X_i X_j \rangle_c, \quad (8)$$

where the second sum is made over the $N_k(N_k - 1)/2$ permutations of i and j without repetitions. If the Fourier modes were statistically independent inside the shell then one could drop the double sum (second term) in Eq. (8). Notice that there is a subtlety that needs to be mentioned when inspecting Eq. (8), despite the fact that we define $X_i := |\delta_{k_i}|^2$ one cannot say that $\langle X_i^n \rangle_c = \langle |\delta_{k_i}|^{2n} \rangle_c$. The equality holds only for the moments but not for the cumulant moments. Indeed, without paying attention to this, looking at Eq. (8) one would erroneously conclude that for a Gaussian field the variance of the shell average vanishes. Instead, one can use the relation between moments $\langle X_i^n \rangle_c = \langle |\delta_{k_i}|^{2n} \rangle_c$ together with the assumption of statistical invariance by translation to show that for one-point statistics

$$\langle X_i^2 \rangle_c = \langle |\delta_{k_i}|^4 \rangle_c + \langle |\delta_{k_i}|^2 \rangle_c^2, \quad (9)$$

¹ We use the following Fourier convention $\delta_k := \frac{1}{(2\pi)^3} \int_{L^3} \delta(\mathbf{x}) e^{-i\mathbf{k}\cdot\mathbf{x}} d^3\mathbf{x}$.

while for two-point statistics

$$\langle X_i X_j \rangle_c = \langle |\delta_{k_i}|^2 |\delta_{k_j}|^2 \rangle_c. \quad (10)$$

As a result, Eq. (8) $\langle X_i X_j \rangle_c$ involves a four-point function of the density field that is null in the Gaussian case. Instead, the first term of Eq. (8) contains an intrinsic contribution that would arise even if the density field were Gaussian (i.e. $\langle |\delta_k|^2 \rangle_c^2$) and a part that would only arise when considering a non-Gaussian density field (i.e. $\langle |\delta_k|^4 \rangle_c$). The usual (see [Scoccimarro et al. 1999](#); [Meiksin & White 1999](#); [Smith 2009](#); [Smith & Marian 2015](#)) expression for the variance of the power spectrum can be obtained by splitting $\langle X_i^2 \rangle_c$ (see Eq. (9)) into its Gaussian and non-Gaussian parts:

$$\langle X^2 \rangle_c = \frac{\bar{P}_\ell^2}{N_k} + k_F^3 \bar{T}_\ell, \quad (11)$$

where we define k as being at the centre of the spherical shell of thickness k_F and V_k its volume. In addition, we define the shell average powers of the power spectrum as

$$\bar{P}_\ell^n := \frac{(2\ell + 1)^n}{V_k} \int_{V_k} P^n(\mathbf{k}_1) \mathcal{L}_\ell^n(\mu_1) d^3 \mathbf{k}_1, \quad (12)$$

where the subscript V_k in the integral means that the integration variables belong to the Fourier shell, k , and thus \bar{P}_ℓ^n depends on the Fourier shell k . We note that we used the continuous limit of the shell average as described in the Appendix A. We also defined the corresponding quantity for the trispectrum as

$$\bar{T}_\ell := \frac{(2\ell + 1)^2}{V_k^2} \int_{V_k} T(\mathbf{k}_1, -\mathbf{k}_1, \mathbf{k}_2) \mathcal{L}_\ell(\mu_1) \mathcal{L}_\ell(\mu_2) d^3 \mathbf{k}_1 d^3 \mathbf{k}_2, \quad (13)$$

which also depends on the considered k -shell.

The same reasoning about moments and cumulant moments can be applied to obtain the formal expression of the skewness:

$$\begin{aligned} \langle X^3 \rangle_c &= \sum_{i=1}^{N_k} a_i^3 \langle X_i^3 \rangle_c + 3 \sum_{j>i} \left[a_i^2 a_j \langle X_i^2 X_j \rangle_c + a_i a_j^2 \langle X_i X_j^2 \rangle_c \right] \\ &+ 6 \sum_{n>j>i} a_i a_j a_n \langle X_i X_j X_n \rangle_c. \end{aligned} \quad (14)$$

Equation (14) allows us to understand that three kinds of correlators can contribute to the skewness, and they are

$$\langle X_i^3 \rangle_c = \langle |\delta_{k_i}|^6 \rangle_c + 6 \langle |\delta_{k_i}|^4 \rangle_c \langle |\delta_{k_i}|^2 \rangle_c + 2 \langle |\delta_{k_i}|^2 \rangle_c^3, \quad (15)$$

$$\langle X_i^2 X_j \rangle_c = \langle |\delta_{k_i}|^4 |\delta_{k_j}|^2 \rangle_c + 2 \langle |\delta_{k_i}|^2 \rangle_c \langle |\delta_{k_i}|^2 |\delta_{k_j}|^2 \rangle_c, \quad (16)$$

$$\begin{aligned} \langle X_i X_j X_n \rangle_c &= \langle |\delta_{k_i}|^2 |\delta_{k_j}|^2 |\delta_{k_n}|^2 \rangle_c + \langle \delta_{k_i} \delta_{k_j} \delta_{k_n} \rangle_c^2 \\ &+ \langle \delta_{-k_i} \delta_{k_j} \delta_{k_n} \rangle_c^2 + \langle \delta_{k_i} \delta_{-k_j} \delta_{k_n} \rangle_c^2 + \langle \delta_{k_i} \delta_{k_j} \delta_{-k_n} \rangle_c^2. \end{aligned} \quad (17)$$

The one-point cumulant (Eq. (15)) is composed of an intrinsic contribution (present even in the Gaussian case) and contributions related to the pentaspectrum and trispectrum. Then, there are two-point correlators (Eq. (16)) and three-point correlators (Eq. (17)) which would both be null in the Gaussian case since they involve only six- and four-point functions. By inserting Eqs. (15)–(17), into Eq. (14), one can express the full expression of the skewness as

$$\langle X^3 \rangle_c = \frac{2}{N_k^2} \bar{P}_\ell^3 + \frac{3}{N_k} k_F^3 \left(\bar{B}_\ell^2 + 2 \bar{M}_\ell \right) + k_F^6 \bar{Q}_\ell, \quad (18)$$

where we define the shell average of the square of the bispectrum as

$$\begin{aligned} \bar{B}_\ell^2 &:= \frac{(2\ell + 1)^3}{V_k^2} \int_{V_k} B^2(\mathbf{k}_1, \mathbf{k}_2) \mathcal{L}_\ell(\mu_1) \mathcal{L}_\ell(\mu_2) \\ &\times \mathcal{L}_\ell(\mu_1 + \mu_2) \Theta_k(k_3) d^3 \mathbf{k}_1 d^3 \mathbf{k}_2; \end{aligned} \quad (19)$$

the shell average of the product between the power spectrum and the trispectrum as

$$\begin{aligned} \bar{M}_\ell &:= \frac{(2\ell + 1)^3}{V_k^2} \int_{V_k} P(\mathbf{k}_1) T(\mathbf{k}_1, -\mathbf{k}_1, \mathbf{k}_2) \\ &\times \mathcal{L}_\ell^2(\mu_1) \mathcal{L}_\ell(\mu_2) d^3 \mathbf{k}_1 d^3 \mathbf{k}_2; \end{aligned} \quad (20)$$

and the shell averaged pentaspectrum as

$$\begin{aligned} \bar{Q}_\ell &:= \frac{(2\ell + 1)^3}{V_k^3} \int_{V_k} Q(\mathbf{k}_1, -\mathbf{k}_1, \mathbf{k}_2, -\mathbf{k}_2, \mathbf{k}_3) \\ &\times \mathcal{L}_\ell(\mu_1) \mathcal{L}_\ell(\mu_2) \mathcal{L}_\ell(\mu_3) d^3 \mathbf{k}_1 d^3 \mathbf{k}_2 d^3 \mathbf{k}_3. \end{aligned} \quad (21)$$

In the present work we aim at quantifying the various terms involved in Eq. (18) and their relative contribution to the reduced skewness,

$$S_3^{(\ell)} := \langle X^3 \rangle_c / \langle X^2 \rangle_c^{3/2}, \quad (22)$$

in various configurations of the density field. In the rest of the paper we refer to the reduced skewness, $S_3^{(\ell)}$, as the skewness, where ℓ refers to the fact that this is the skewness of the multipole ℓ of the power spectrum², and we refer to $\langle X^3 \rangle_c$ as the third-order cumulant.

3. Simulations and estimators

In this section we first present the different sets of simulations we generated and analysed in this study. Then, we present the different estimators employed to probe the distribution of the estimated power spectra.

3.1. Approximate simulations with COVMOS

The standard N -body approach is a widely used technique for simulating the evolution of large-scale cosmic structures. In our case, however, the computational cost of producing the large number of realisations required for a robust analysis of the power spectrum distribution is prohibitive. This is why in this study we exclusively rely on an approximate mock-catalogue generator, COVMOS³ ([Baratta et al. 2020, 2023](#)).

The COVMOS public code enables rapid simulation of catalogues for various cosmological objects in both real and redshift space. Its speed is due to the fact that it does not require the evolution of an initial density, as typically needed in an N -body process. Instead, a density field is directly generated at a specific redshift and then subjected to local Poisson sampling to create the point-like catalogue.

Although similar approaches, such as log-normal mock generators, are extensively studied in the literature ([Xavier et al. 2016](#); [Alonso et al. 2014](#); [Agrawal et al. 2017](#)), the COVMOS method has two distinctive features. First, the power spectrum of

² Note that this is an abuse of notation because $S_3^{(\ell)}$ is not actually a multipole of the Legendre expansion as $P^{(\ell)}(k)$.

³ github.com/PhilippeBaratta/COVMOS

Table 1. Redshift at which the COVMOS mock catalogue samples considered in this work have been generated.

Name	z
Λ CDM	0, 0.5, 1, 1.5, 2
16nu	0
w9p3	0

the density and velocity fields, as well as the density one-point PDF, can be arbitrarily set by the user. This flexibility means it is not restricted to a log-normal PDF, enabling the exploration of more realistic models. These statistics can be provided either from analytical models or directly from estimates based on simulations or observational data.

The second distinctive feature of the COVMOS method is its ability to assign specific velocities to objects based on the local density, while ensuring the validity of the continuity equation. This capability leads to a high simulation performance for RSD effects on two-point statistics, covering both the linear and mildly non-linear regimes.

The COVMOS method relies on the application of a local transform to a Gaussian field in order to match a target one-point PDF. As discussed in Baratta et al. (2023), this means that the higher order one-point moments of the PDF are correctly reproduced. Since they are given by specific integrals of the Fourier space N -point correlation functions, it is expected that COVMOS is able to capture some of the higher order information. Indeed, in Baratta et al. (2023) it has been shown that the non-Gaussian part of the covariance can be well recovered at least up to $k \sim 0.25 h/\text{Mpc}$ at $z = 0$. In addition, as shown in Sect. 2, the skewness of the distribution of the power spectrum depends on higher order correlation functions. This is the reason why we designed a validation procedure in Appendix B that justifies the use of the COVMOS method for this study.

In this study, we generated numerous large samples of dark matter catalogues in both real and redshift-space, all within simulated periodic boxes of volume $1 (\text{Gpc}/h)^3$. However, a cut in geometry of these catalogues, involving survey masks, is performed in Sect. 4.3.2. These samples are detailed in Table 1, which includes the sample name, simulated cosmology, and redshifts. Each sample contains 100 000 realisations of 10^8 particles, thus a density of $0.1 (\text{Mpc}/h)^{-3}$, this means that the shot-noise power spectrum is equal to the input power spectrum at $k = 1.65 h/\text{Mpc}$.

We selected three different flat cosmologies compatible with the Planck 2013 results (Planck Collaboration XVI 2014), with cosmological parameters $(\Omega_m, \Omega_b, h, n_s, A_s) = (0.32, 0.05, 0.67, 0.96, 2.1265 \times 10^{-9})$. The first cosmology is characterised by a Lambda cold dark matter (Λ CDM) background, the second (16nu) includes massive neutrinos with total mass of $\sum m_\nu = 0.16 \text{ eV}$. In the third cosmology, named w9p3, a dynamical dark energy is included using the CPL (Chevallier & Polarski 2001; Linder 2003) parametrisation, where $w(a) = w_0 + w_a(1 - a)$ and $(w_0, w_a) = (-0.9, 0.3)$.

As previously stated, to use the COVMOS method, the user must specify target values for the power spectrum and one-point PDF of the density field, along with the target velocity power spectrum. For all our samples, we adopted these target values based on measurements performed in the DEMNUni suite of N -body simulations (Carbone et al. 2016) which have been run for the cosmological models described above. In the case

of the Λ CDM and 16nu cosmology we use the one-point PDF and power spectrum, averaged over the 50 DEMNUni-Cov realisations (Baratta et al. 2023; Gouyou Beauchamps et al. 2025; Paribelli et al. 2021), with a sampling of 512 points per side. For the w9p3 cosmology we use the unique realisation of this cosmological model from the DEMNUni set of simulations (Paribelli et al. 2022) with a sampling of 1024 points per side (comoving box twice larger).

3.2. Statistics estimators

To estimate the power spectrum, we employed the widely used NBodyKit⁴ software (Hand et al. 2018), which offers robust and efficient tools for analysing the multipoles of the power spectrum, both in real and redshift space. In particular it allows the user to vary different settings of the estimation procedure, including the type of mass assignment scheme employed to interpolate particles on a grid, or the activation or not of the interlacing technique allowing the user to reduce the aliasing effect. We take advantage of this opportunity to identify a possible impact of these settings on the distribution of the output power spectra.

Despite having introduced (in Sect. 2) only the lowest level of non-Gaussianity (i.e. the skewness) of the distribution of the measured power spectrum, in the present work we also estimate the reduced kurtosis $S_4^{(\ell)} := \langle X^4 \rangle_c / \langle X^2 \rangle_c^2$. We apply the same nomenclature to the kurtosis, i.e. $S_4^{(\ell)}$ is dubbed kurtosis, and $\langle X^4 \rangle_c$ is dubbed the fourth-order cumulant.

The estimation of the skewness and kurtosis is based on the $N_s = 100\,000$ realisations that we generated with the COVMOS method, and their estimators are

$$\hat{S}_3^{(\ell)}(k) = \frac{N_s^{-1} \sum_{i=1}^{N_s} [\hat{P}_i^{(\ell)}(k) - \bar{P}^{(\ell)}(k)]^3}{\left\{ N_s^{-1} \sum_{i=1}^{N_s} [\hat{P}_i^{(\ell)}(k) - \bar{P}^{(\ell)}(k)]^2 \right\}^{3/2}} \quad (23)$$

and

$$\hat{S}_4^{(\ell)}(k) = \frac{N_s^{-1} \sum_{i=1}^{N_s} [\hat{P}_i^{(\ell)}(k) - \bar{P}^{(\ell)}(k)]^4}{\left\{ N_s^{-1} \sum_{i=1}^{N_s} [\hat{P}_i^{(\ell)}(k) - \bar{P}^{(\ell)}(k)]^2 \right\}^2} - 3, \quad (24)$$

where $\hat{P}_i^{(\ell)}(k)$ is our estimate of the ℓ -th multipole of the power spectrum in shell k for the i -th COVMOS realisation, and $\bar{P}^{(\ell)}(k)$ is the estimated average multipole power spectrum over the N_s COVMOS realisations,

$$\bar{P}^{(\ell)}(k) := \frac{1}{N_s} \sum_{i=1}^{N_s} \hat{P}_i^{(\ell)}(k). \quad (25)$$

Note that since we are estimating the mean $\bar{P}^{(\ell)}(k)$, the estimators of the skewness and kurtosis are expected to be biased. However, given the extremely high number of realisations, N_s , we can safely neglect that bias. Indeed, the variance would be affected at the level of $N_s/(N_s - 1)$, and the skewness would be affected at the level of $N_s^2/(N_s - 2)/(N_s - 1)$ (i.e. adjusted Fisher-Pearson standardised moment coefficient), which is at most of the order of 0.3%.

The estimation of the skewness and kurtosis should be compared to the corresponding prediction assuming a Gaussian density field to quantify their excess. In Appendix A we review

⁴ nbodykit.readthedocs.io

the general Gaussian field prediction for $S_n^{(\ell)}$, which if one further assumes an isotropic real space power spectrum, it can be expressed as

$$S_{3,G}^{(0)}(k) = \frac{2}{\sqrt{N_k}}, \quad (26)$$

$$S_{4,G}^{(0)}(k) = \frac{6}{N_k}, \quad (27)$$

for the monopole power spectrum estimator, where N_k is the number of independent wave modes in the considered k -shell. One can notice that Eqs. (26) and (27) apparently differ from Takahashi et al. (2009), but this is only due to a different definition of the number of modes, N_k . Indeed, they defined N_k as the total number of modes, while we restrict it to the independent modes (as previously explained). Thus, N_k can either be computed exactly on a Fourier grid or approximately evaluated analytically by dividing the total Fourier volume of the shell ($2\pi k^2 k_F$) by the Fourier volume (k_F^3) of an individual mode, yielding

$$N_k \simeq 2\pi (k/k_F)^2. \quad (28)$$

Note that Eqs. (26) and (27) are different when considering the quadrupole and the hexadecapole. We show in Appendix A that

$$S_{3,G}^{(2)}(k) = \sqrt{5} \frac{2}{7} \frac{2}{\sqrt{N_k}}, \quad (29)$$

$$S_{4,G}^{(2)}(k) = \frac{15}{7} \frac{6}{N_k}, \quad (30)$$

$$S_{3,G}^{(4)}(k) = 3 \frac{162}{1001} \frac{2}{\sqrt{N_k}}, \quad (31)$$

$$S_{4,G}^{(4)}(k) = 2.5180114 \frac{6}{N_k}. \quad (32)$$

In redshift space the expressions are more complicated but knowing the multipoles of the power spectrum one can still evaluate the corresponding expressions numerically on the 3D Fourier grid as given by Eq. (A.10) in Appendix A.

4. Quantifying the non-Gaussianity of the distribution of the estimated power spectrum

In this section we present a phenomenological study of the power spectrum distribution when varying different parameters. First, we study the cosmological effects that may impact the various statistics due to non-linear clustering. This includes the cosmological model and the redshift. Then, we investigate the effect of the power spectrum estimator's settings through the Fourier mode binning and the NBodyKit input parameters for the 3D Fourier grid. Finally, we look at survey specific features, namely the sample density and the survey mask.

4.1. Cosmological effects

In this subsection we want to explore the potential cosmological factors that could influence the distribution of the power spectrum estimator, primarily as a result of non-linear clustering. Thanks to the analytical framework developed in Sect. 2 we are able to quantify the different terms contributing to the non-Gaussianity of the distribution.

In Fig. 1 we present the estimated skewness $S_3^{(\ell)}$ and kurtosis $S_4^{(\ell)}$ of the distribution of the multipoles of the power spectrum

obtained from the sets of Λ CDM simulations at five different redshifts ranging from $z = 0$ to $z = 2$ in real space. The fact that we do not include redshift space distortions allowed us to compare these measurements to the Gaussian predictions of the skewness $S_3^{(\ell)}$ and kurtosis $S_4^{(\ell)}$ (see Eqs. (26)–(32)). One can see that on large scales (low k), both the skewness and kurtosis of the distributions are systematically higher than zero, which agrees with the expectation for a Gaussian field (see Eqs. (26)–(32)). Indeed, as the wave number k decreases, the number of wave modes, N_k , reduces, thus significantly increasing the skewness and kurtosis.

Inspecting the left panel of Fig. 1, we can see that only the high- k part of the monopole's skewness shows significant deviation with respect to the Gaussian field prediction. This is in good agreement with Blot et al. (2015) and with Takahashi et al. (2009). In addition, we can see that this deviation is more important at a low redshift. At $z = 0$ we can see the excess of skewness starting at $k = 0.1 h/\text{Mpc}$, while at $z = 1$ it starts at $k = 0.2 h/\text{Mpc}$, and we cannot even see it at $z = 2$ for the scales probed here. This behaviour is expected as the terms \bar{B} , \bar{T} , and \bar{Q} in Eq. (18) are becoming more important through the non-linear evolution of the matter clustering. Note that when k increases, the skewness starts to decrease again but less rapidly than in the Gaussian case; thus one could think that in the limit of an infinite number of modes the skewness is expected to vanish, which is in agreement with the central limit theorem.

Regarding the kurtosis (right panel of Fig. 1) it is more difficult to draw strong conclusions because the estimation is more affected by noise. Focusing on the monopole we can detect some excess of kurtosis around $k > 0.3 h/\text{Mpc}$ at $z = 0$. However, for $z > 0.5$ there is no direct detection of an excess kurtosis.

In redshift space (see Fig. 2), the excess of skewness observed at high wave modes for the monopole in real space is suppressed, yielding a better agreement with the Gaussian prediction. The excess skewness of the monopole in real space appears to be transferred to the quadrupole and hexadecapole in redshift space, causing deviations from their respective predictions. Specifically, the skewness of the quadrupole decreases and becomes negative, while the skewness of the hexadecapole increases. Regarding the kurtosis, no excess can be detected at any scale, redshift, and multipole.

Since it appears that the dominant non-Gaussian contribution is the skewness, in the following we put the kurtosis aside and quantify the contribution of each term in Eq. (18) to the total skewness.

We define the relative difference of the total cumulant moments with respect to their Gaussian only contribution as

$$r_i := \frac{\langle X^i \rangle_c}{\langle X^i \rangle_G} - 1. \quad (33)$$

From Eq. (11), one can show that r_2 , which is the non-Gaussian relative contribution to the variance, is given by

$$r_2 = N_k k_F^3 \frac{\bar{T}_0}{\bar{P}_0^2}. \quad (34)$$

It is directly related to the shell averaged trispectrum \bar{T}_0 .

Based on the structure of Eq. (18), we can write r_3 as

$$r_3 = \frac{3}{2} N_k k_F^3 \frac{\bar{B}_0^2 + 2\bar{M}_0}{\bar{P}_0^3} + \frac{N_k^2 k_F^6}{2} \frac{\bar{Q}_0}{\bar{P}_0^3}. \quad (35)$$

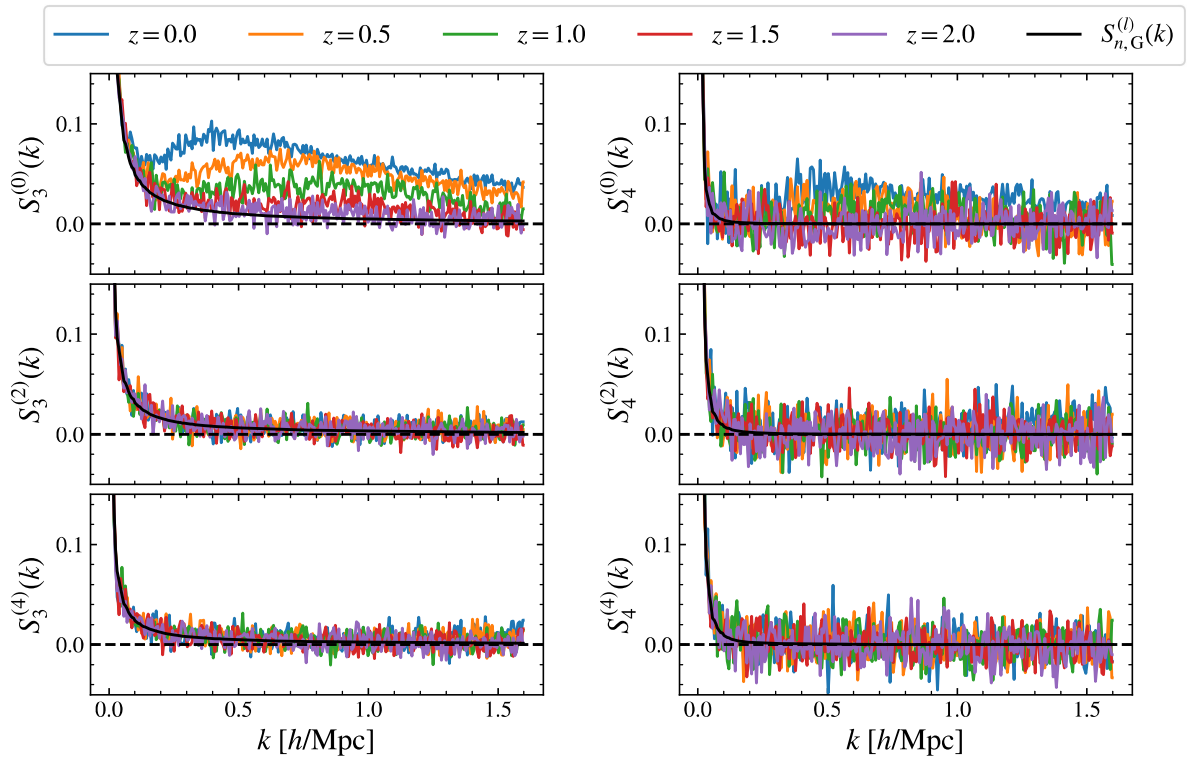


Fig. 1. Estimated skewness, $S_3^{(\ell)}$ (left panels), and kurtosis, $S_4^{(\ell)}$ (right panels), of the distribution of power spectrum multipoles $P^{(\ell)}(k)$ in real space for the reference Λ CDM cosmology and for the five redshifts considered in this work (from top to bottom panels $\ell = 0, 2, 4$). In each panel, the black line shows the prediction for a Gaussian field.

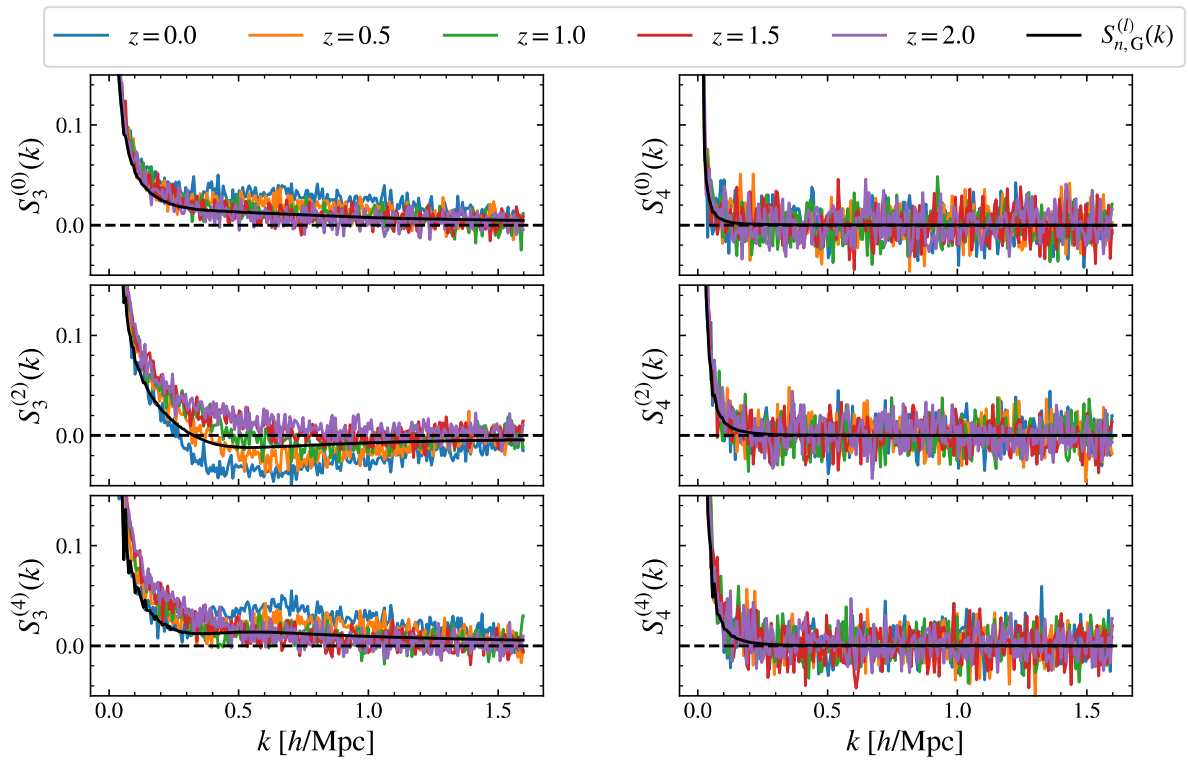


Fig. 2. Same as Fig. 1 but in redshift space.

When considering the monopole in real space, one can show that $\bar{M}_0 = P_0 \bar{T}_0$ and $\bar{P}_0^n = P_0^n$, which allows us to rewrite r_3 as

$$r_3 = b + 3r_2 + \frac{N_k^2 k_F^6 \bar{Q}_0}{2 P_0^3}, \quad (36)$$

where

$$b := \frac{3}{2} N_k k_F^3 \bar{B}_0^2. \quad (37)$$

Equation (36) highlights the fact that different terms contribute to the non-Gaussian part of $\langle X^3 \rangle_c$, specifically: the bispectrum term b , the trispectrum term $3r_2$, and the pentaspectrum term.

We start by showing that the bispectrum contribution, b , is expected to be negligible with respect to $3r_2$. For this we can notice two things about the three modes \mathbf{k}_1 , \mathbf{k}_2 , and \mathbf{k}_3 (see Eq. (3)) defining a bispectrum configuration: first, invariance by translation imposes that $-\mathbf{k}_3 = \mathbf{k}_1 + \mathbf{k}_2$, and second, we need their modulus to be contained within the considered k -shell. These two constraints mean that not all triplets, \mathbf{k}_1 , \mathbf{k}_2 , \mathbf{k}_3 , will contribute to the shell average of the bispectrum. Indeed, only pairs \mathbf{k}_1 , \mathbf{k}_2 that satisfy $|\mathbf{k}_1 + \mathbf{k}_2| \approx k$ can contribute. Thus, only nearly equilateral configurations of the triplets will be kept in the shell averaging. We can see in Eq. (17) that there are four kinds of configurations for the bispectrum $\mathbf{k}_1 + \mathbf{k}_2$, $\mathbf{k}_1 - \mathbf{k}_2$, $-\mathbf{k}_1 - \mathbf{k}_2$ and $\mathbf{k}_2 - \mathbf{k}_1$ with respectively N_1 , N_2 , N_3 and N_4 number of triangular configurations which are contributing to the shell-averaged square of the bispectrum \bar{B}_0^2 . As a result, defining $N = N_1 + N_2 + N_3 + N_4$ as the total number of triangles that contribute to the shell average and using symmetry conditions one can show that the total (i.e. double counting the repeated conjugate modes) number of triangles is roughly given by $N \approx \pi^2 (k/k_F)^3$. In addition, from Eq. (19) one can show that the monopole \bar{B}_0 is related to the equilateral configuration $B(k)$ of the bispectrum through

$$\bar{B}_0^2 \approx \frac{N}{N_k^2} B^2(k). \quad (38)$$

In practice, when considering only half of the Fourier space (discarding non-independent modes) the four configurations of the triangles formed by \mathbf{k}_1 and \mathbf{k}_2 are not appearing with the same rate. This can be estimated directly on a Fourier grid and is represented in Fig. 3. In this figure we show that the total number N of non-degenerate triangular configurations is indeed following the expected trend. One can immediately understand that since $N \propto k^3$ and $N_k^2 \propto k^4$ (see Eq. (28)) the ratio will lead to a suppression of the equilateral configuration of the bispectrum proportional to k^{-1} . We thus expect that the term b is counting less than the trispectrum contribution, which is not suffering from this mode selection within the shell. In fact, assuming that $B(k) \approx P^3(k)$ and that $\bar{T}(k) \approx P^4(k)$ (see Scoccimarro et al. 1999; Baratta et al. 2020, for a justification), one can show that we expect

$$b \approx 3r_2 \frac{N}{N_k^2}. \quad (39)$$

By estimating the second- and third-order cumulants from the Λ CDM simulation at $z = 0$, we can compute $3r_2$ and r_3 . In addition, we can estimate the equilateral configurations of the bispectrum $B(k)$ on 100 000 realisations. It allows us to evaluate the contribution b to the relative excess skewness r_3 . In Fig. 4 we compare these three quantities. First, we can see that the overall amplitude of the bispectrum term b is indeed of the order

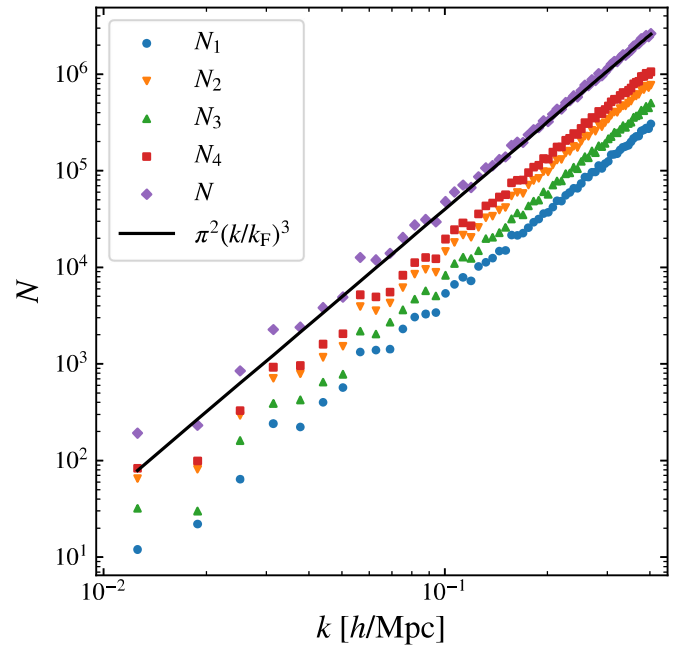


Fig. 3. Number of triplets, \mathbf{k}_1 , \mathbf{k}_2 , \mathbf{k}_3 , per k -shell depending on the considered configuration, N_1 , N_2 , N_3 , and N_4 . The purple diamonds show the total number of triplets, and the black line shows the corresponding expected number.

expected from Eq. (39), which is displayed as a gray dotted line. Second, due to the suppression mechanism described above, the b contribution is more than one order of magnitude smaller than the trispectrum contribution $3r_2$ at the lowest accessible k and more than two orders of magnitude smaller at $k = 0.2 h/\text{Mpc}$. This justifies neglecting the bispectrum term b in the following sections. Thus, we are left with two terms that contribute to the relative excess of skewness r_3 : the trispectrum term $3r_2$ and the pentaspectrum term which, neglecting b , is simply $r_3 - 3r_2$.

In Fig. 5 we show how the non-Gaussian contribution to the covariance evolves with redshift in the COVMOS realisations. In this figure $3r_2$ is represented in dashed line so when it crosses the level 3 (dotted line) it means that the Gaussian and the non-Gaussian part contribute the same amount to the total variance. This is confirming that when the redshift is low ($z = 0$), the term $N_k k_F^3 \bar{T}_0 / \bar{P}_0^2$ starts dominating the variance already at $k = 0.25 h/\text{Mpc}$ while at high redshift ($z = 2$) it happens on smaller scales ($k = 0.6 h/\text{Mpc}$). This is a well known result which was already shown by Takahashi et al. (2009).

In addition, in Fig. 5 we show the contribution of the two terms $r_3 - 3r_2$ (pentaspectrum) and $3r_2$ (trispectrum), to r_3 (excess skewness) as a function of k for the five redshifts. It appears that at low redshift the term $r_3 - 3r_2$ starts dominating already at $k > 0.3 h/\text{Mpc}$. However, at high redshift this term fluctuates around 0 and r_3 is rather dominated by $3r_2$.

Based on the above observations and by expressing the ratio between the total skewness and its prediction for a Gaussian field in terms of r_2 and r_3 as

$$\frac{S_3^{(0)}}{S_{3,G}^{(0)}} = \frac{1 + r_3}{(1 + r_2)^{3/2}}, \quad (40)$$

we can obtain approximate expressions in the two regimes (either the pentaspectrum or the trispectrum domination regime) described above. At high redshift, when r_3 is dominated by $3r_2$

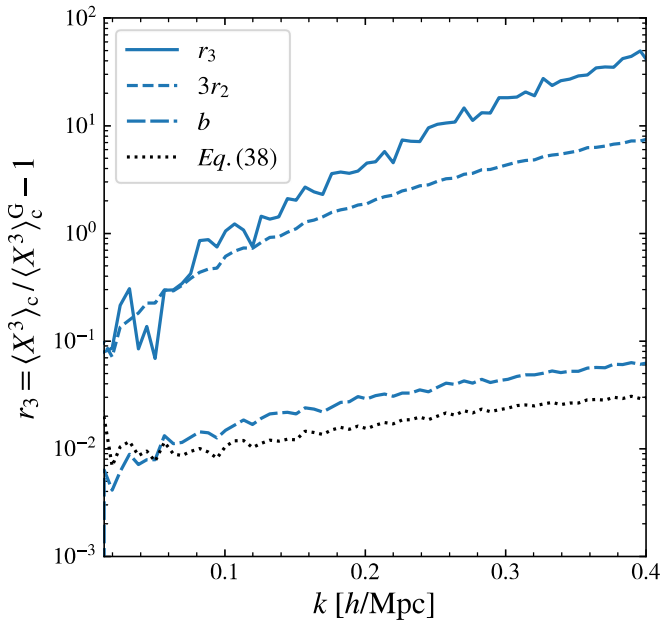


Fig. 4. Measurement of r_3 from the Λ CDM simulation at $z = 0$ and its different contributions. The solid line shows the total relative difference r_3 , while the dashed and dot-dashed lines show the trispectrum and bispectrum contribution, $3r_2$ and b , respectively. The dotted black line shows a rough estimation of b based on Eq. (39).

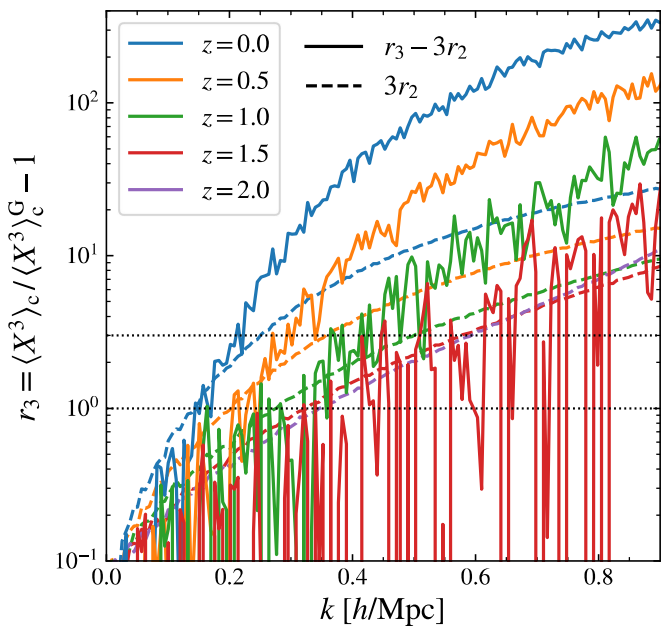


Fig. 5. Relative excess skewness, r_3 , made of two contributions: $3r_2$ and $r_3 - 3r_2$ (respectively in dashed and solid lines). The colours correspond to the different redshifts. For clarity of the figure, we do not show $r_3 - 3r_2$ for $z = 2$, as it is compatible with 0. There are two horizontal dotted black lines showing the levels 1 and 3.

(trispectrum domination), a simple calculation shows that when $3r_2 \approx 9$ the ratio in Eq. (40) deviates from 1 by only 25%. On the other hand, when we consider low redshift and relatively non-

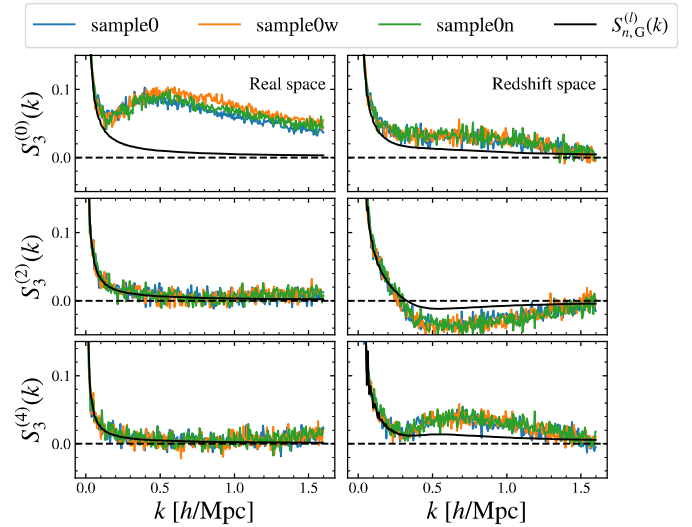


Fig. 6. Estimated skewness $S_3^{(\ell)}$ of the distribution of the power spectrum multipoles (from top to bottom panels $\ell = 0, 2, 4$) $P^{(\ell)}(k)$ in real (left) and redshift space (right) for the Λ CDM cosmology and its two extensions, 16nu and w9p3. In each panel the black line is the skewness expected for a Gaussian field.

linear scales ($k > 0.3 h/\text{Mpc}$) we can use the approximation

$$r_3 \approx \frac{N_k^2 k_F^6 \bar{Q}_0}{2 P_0^3}. \quad (41)$$

As a result, in that regime one can approximate Eq. (40) as

$$\frac{S_3^{(0)}}{S_{3,G}^{(0)}} \approx \frac{\sqrt{N_k k_F^3} \bar{Q}_0}{2 \bar{T}_0^{3/2}}. \quad (42)$$

This shows that the excess of skewness is directly related to the ratio between the shell averaged pentaspectrum and the shell averaged trispectrum to the power of 3/2. As a result, Eq. (42) shows that if this ratio is not well reproduced in the COVMOS realisations we would not get a realistic excess of skewness.

Thus, one might wonder about the soundness of the COVMOS method to study the non-Gaussian behaviour of the distribution of the power spectrum estimator. To gain confidence in the COVMOS method in terms of higher-order statistics, we conducted a comparison between the COVMOS Λ CDM simulation at $z = 0$ and a set of 12 288 N -body simulations using a comparable cosmology, the DEUS-PUR simulations (Blot et al. 2015). As demonstrated in Appendix B, our results using COVMOS exhibit a robust agreement with those obtained from N -body simulations, up to $k = 0.8 h/\text{Mpc}$. This comparison further validates the robustness and accuracy of the COVMOS method for analysing non-Gaussian features.

Since the skewness $S_3^{(\ell)}$ of the estimator of the multipoles of the power spectrum depends explicitly on higher-order correlation functions of the density field we also investigate how it changes when considering different cosmological models characterised by roughly the same power spectrum. In Fig. 6 we compare the skewness estimated in the three samples, Λ CDM, 16nu and w9p3, corresponding to three different cosmological models (see Table 1) at $z = 0$, both in real and redshift space. We can see that adding massive neutrinos (with 16nu) is not changing the skewness, while modifying the dark energy equation of state is only slightly affecting the large k behaviour of the real

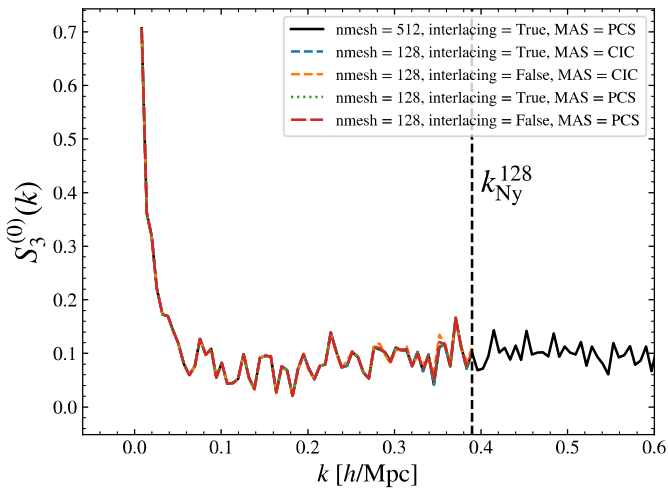


Fig. 7. Estimated skewness of the power spectrum monopole in real space for different estimator settings on a reduced set of 10 000 realisations.

space monopole. In redshift space all multipoles skewness are unchanged by varying cosmology. It is worth noticing that all three models exhibit similar clustering amplitudes, as evidenced by their power spectra presented in Parimbelli et al. (2022).

4.2. Effects of the estimator's setting on the data distribution

In this subsection we investigate how technical details of the power spectrum estimator can impact its statistical distribution. We focus on the real-space distribution, which maximises the excess of skewness.

4.2.1. Aliasing and mass assignment scheme

Despite the soundness of the results presented above, we want to verify their robustness against changes in the choices made for the settings of the estimator. One could think of a possible dependence in the choice of the mass assignment through aliasing effects. Indeed, aliasing is introducing some correlations between modes which might induce some unexpected skewness (Baratta et al. 2020). Thus, to ensure that the observed excess of skewness in the power spectrum distribution is indeed cosmological and do not arise from any numerical artefact, we modified and compared different settings of the estimator.

We tested three different settings. The first was (i) the interlacing technique (Sefusatti et al. 2016), in which one can choose to reduce aliasing arising when using fast Fourier transforms. The second concerned (ii) the type of mass assignment scheme used to interpolate the particles over the grid before estimating the power spectrum, where we tested two of them: the cloud-in-cell and piecewise cubic spline. The third involved (iii) the precision of the grid, where we chose a standard mesh of $N_{\text{mesh}} = 512$ points per dimension and a less precise one with $N_{\text{mesh}} = 128$.

The results are shown in Fig. 7. They clearly demonstrate that, regardless of the choice of these three settings, the skewness of the power spectrum distribution remains the same, including when close to the Nyquist mode $k_{\text{Ny}}^{N_{\text{mesh}}} = \pi N_{\text{mesh}}/L$. This indicates that the type of estimator used does not affect the skewness of the power spectrum.

4.2.2. Fourier mode binning

In galaxy survey analyses, it is common to adapt the power spectrum wave-mode bin width Δk to accurately capture the cosmological information across various scales. The goal is to find the optimal Δk to balance the need for detailed understanding of the power spectrum behaviour across scales while minimising the noise, the correlations between different bins and the number of bins in the data-vector. For example it is possible to enlarge the binning in k such that one can reduce the size of the data-vector. This is particularly useful to obtain cosmological constraints based on a covariance matrix which is estimated from a finite set of realisations (Gouyou Beauchamps et al. 2025).

We thus want to observe the effect of the bin width on the skewness of the estimator of the power spectrum distribution in light of the analytical formalism presented in Sect. 2. Indeed, we expect that enlarging the bin width will affect the shell averaged poly-spectra entering in the third-order cumulant of the distribution of the power spectrum estimator (see Eq. (18)). In the limit of a Gaussian density field, where the skewness is given by Eq. (26), the only effect we can expect is that, by increasing the size of the bins, the number of modes within the Fourier shell will also increase, resulting in a decrease of the skewness. However, considering a non-Gaussian density field it is difficult to predict the effect of the binning with simple arguments. As a result, in this section we systematically test the effect of increasing the bin width of the estimated monopole power spectrum in real space.

Figure 8 displays the excess skewness of the power spectrum distribution with respect to the Gaussian field prediction, at $z = 0$ and in real space, for various bin widths, ranging from $\Delta k = k_{\text{F}} = 0.006 h/\text{Mpc}$ to $\Delta k = 0.1 h/\text{Mpc}$. We can see that increasing the bin width produces more excess skewness, especially on intermediate scales, between $k = 0.15$ and $0.4 h/\text{Mpc}$. In that range we can see that the excess skewness can increase by 66% going from the reference binning size (i.e. the fundamental frequency) to largest one. On these scales higher-order statistics, and thus the non-Gaussian terms in Eq. (18), start to be relevant. On the contrary, for $k > 0.55 h/\text{Mpc}$ the increase with the binning size is less than 15%. Since in this regime the skewness can be approximated with Eq. (42), it means that the ratio between the pentaspectrum and trispectrum to the power $3/2$ is close to be constant when increasing the shell size.

In the lower panel of Fig. 8 we can appreciate how the skewness is increasing with the binning size at fixed $k = 0.3 h/\text{Mpc}$. It appears that it saturates at around 7 times the fundamental frequency. In conclusion, even a drastic increase in the binning size does not increase the skewness of the estimator of the power spectrum more than 50%.

4.3. Effects of survey-related features on the data distribution

In this subsection, we investigate the impact of characteristic features of a survey on the statistical distribution of the power spectrum estimator. Since we have shown in Sect. 4 that being in real space maximises the excess of skewness we keep the same setting in the following.

4.3.1. Sample density

First, we examine the effect of the number density of objects in the catalogue. Indeed, since shot noise affects all N -point correlation functions, it should induce a non-zero trispectrum and pentaspectrum and can thus modulate the expected skew-

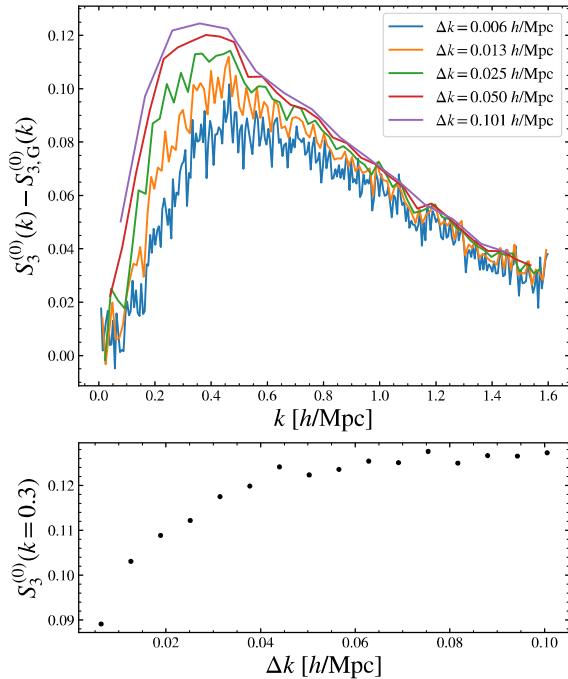


Fig. 8. Impact of wave mode binning. *Top*: Difference between the estimated skewness and the Gaussian field prediction for the real-space power spectrum at $z = 0$ for different bin sizes, Δk . *Bottom*: Skewness at $k = 0.3 h/\text{Mpc}$ with respect to Δk .

ness when the Gaussian contribution is sub-dominant. Our reference catalogue, ΛCDM at $z = 0$, has a number density $\rho = 0.1 (h/\text{Mpc})^3$. We generated a set of three sparser catalogues with densities of $\rho = 0.05, 0.01, \text{ and } 0.001 (h/\text{Mpc})^3$ and for each of them we estimate the skewness of the real-space monopole power spectrum.

The top panel of Fig. 9 displays the average power spectrum of the reference sample, as well as the shot noise amplitude of the other samples. The middle panel shows the estimated skewness $S_3^{(0)}(k)$ in all samples. From this figure, it is evident that increasing the shot noise level tends to lower the excess skewness. The suppression starts to be effective at the wave mode for which the power spectrum and the shot noise level have the same amplitude. One can notice that when shot noise dominates completely over the signal, no excess skewness becomes detectable. This is in agreement with the expectation. In fact, when the signal is dominated by shot noise we expect the power spectrum to be equal to $P_{\text{SN}} = 1/\rho$. In addition, the higher-order N -point spectra such as the bispectrum, trispectrum, and pentaspectrum are affected in a non-trivial way by the shot noise (see Blot et al. 2015, for the explicit expressions up to the trispectrum). However, it is fairly intuitive that a pure Poisson distribution will be characterised by $B(\mathbf{k}_1, \mathbf{k}_2) = P_{\text{SN}}^2$, $T(\mathbf{k}_1, \mathbf{k}_2, \mathbf{k}_3) = P_{\text{SN}}^3$ and $Q(\mathbf{k}_1, \mathbf{k}_2, \mathbf{k}_3, \mathbf{k}_4, \mathbf{k}_5) = P_{\text{SN}}^5$. Thus, at high k we expect $\langle X^2 \rangle_c \simeq k_F^3 P_{\text{SN}}^3$ and $\langle X^3 \rangle_c \simeq k_F^6 P_{\text{SN}}^5$, from which it follows that

$$S_3 \simeq (k_F^3 P_{\text{SN}})^{1/2}, \quad (43)$$

when the signal is dominated by shot noise. The Eq. (43) suggests that the skewness should increase with the shot-noise power spectrum P_{SN} which seems contrary to what is observed in Fig. 9. However, considering the most extreme shot-noise case ($\rho = 0.001 (h/\text{Mpc})^3$) we expect the skewness to decrease to a

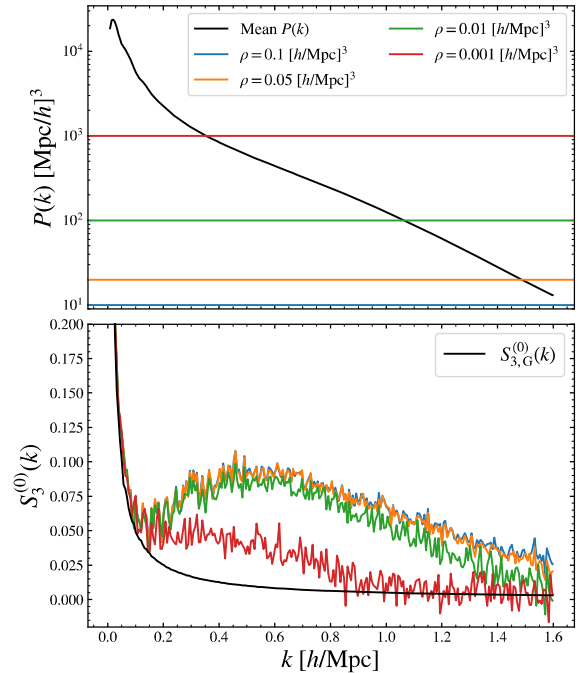


Fig. 9. Impact of the shot noise. *Top*: Mean power spectrum for the reference sample and the different level of shot noise corresponding to each density. *Bottom*: Estimated skewness for the different densities considered (coloured lines) and the Gaussian field prediction (black line).

value of 0.016 at large k , which is a level too low to be detected in our case. The fact that the shot noise reduces the skewness was also observed in Lin et al. (2020) in the case of weak lensing two-point correlation functions.

In Euclid Collaboration: Mellier et al. (2025) it is predicted that the total spectroscopic sample of *Euclid* will have a density of less than $\rho = 10^{-3} (h/\text{Mpc})^3$, which is comparable to the highest level of shot noise shown in Fig. 9.

4.3.2. Survey window function

As observed in Upham et al. (2021), another observational aspect that can significantly impact the distribution of two-point statistics is the survey mask. In order to investigate this effect, we carve two arbitrary cone-like geometries within the periodic boxes generated with COVMOS. We define two different cones that both have an aperture of semi-angle $\theta = \pi/9$. The big cone has a radial extension of $[r_{\text{min}}, r_{\text{max}}] = [600, 1400] \text{ Mpc}/h$ and the small one $[r_{\text{min}}, r_{\text{max}}] = [800, 1000] \text{ Mpc}/h$. The big and small cones have respectively a volume of $0.3 (\text{Gpc}/h)^3$ and $0.06 (\text{Gpc}/h)^3$. The particle number density is the same as the reference value of $\rho = 0.1 (h/\text{Mpc})^3$ that we considered for the periodic box. In Fig. 10 we show how the small and big cones are compared with the reference comoving box.

We estimate the power spectrum in each of the 100 000 realisations at $z = 0$, in real space, for both cone-like geometries with an enclosing box of sided $1000 h^{-1} \text{ Mpc}$. As we are no longer analysing a periodic box we need to take care about the effect of the mask in the estimation of the power spectrum (see Beutler et al. 2017 or GB-2025 for more details). To have an accurate estimation of the survey mask power spectrum we take the average power spectrum from 50 realisations of a uniform random distribution following the cone geometries with 20 times the reference density.

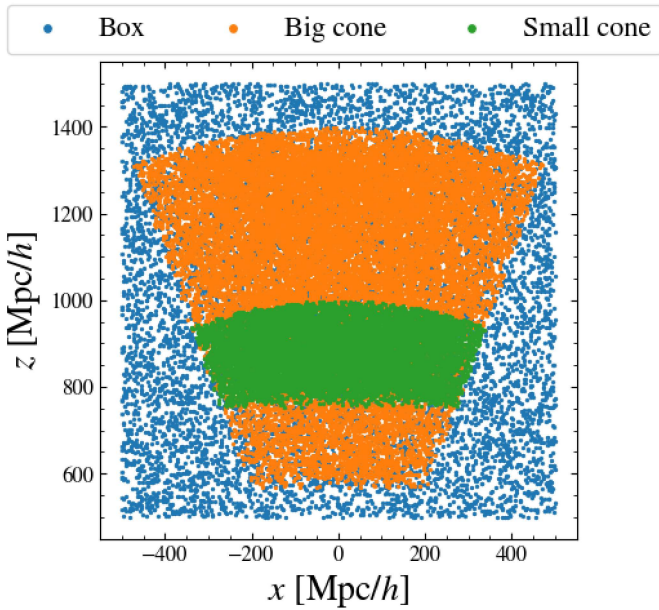


Fig. 10. Two-dimensional projection of a sample of particles from the three different geometries considered: the periodic box and the big and small cone.

From the distribution of the measured power spectra we can estimate the skewness of its distribution (Eq. (23)) in the small and big cones and compare them to the reference periodic box. In addition, we compare these to the Gaussian field prediction. However, if in the case of a periodic box it is fairly straightforward to get this prediction (see Appendix A), in the case of a masked field calculations become significantly more complicated. We need to be careful about the effect of the integral constraint (IC, de Mattia & Ruhlmann-Kleider 2019), which is expected to propagate into the skewness.

The IC comes from the fact that the density contrast is defined from the mean number density of objects in a given catalogue. However, we can only estimate it from the same volume that we have at our disposal to estimate two-point statistics. This gives rise to correlation between large- and short-scale modes. Thus, in order to properly compare the skewness estimated in the COVMOS realisations with respect to the Gaussian field case, we generate $N_G = 100\,000$ realisations of masked Gaussian fields with and without taking into account the IC. We set-up these Gaussian fields, dubbed $\nu(\mathbf{x})$ in the following, such that they have the same power spectrum as the COVMOS realisations.

In the following we explain how we include or not the contribution from IC in these Gaussian realisations. As explained above, the mean number density \bar{n} that we estimate inside a given mask of volume V , as $\bar{n} = N_{\text{gal}}/V$, is different from the true number density $n_0 = \rho$. Thus, within the same volume V the estimated density contrast $\delta_{\text{IC}}(\mathbf{x})$ is different from the true one $\delta(\mathbf{x})$. The relation between the two can be obtained by noticing that the only quantity which is conserved is the local number density $n(\mathbf{x})$, implying that

$$\bar{n} [1 + \delta_{\text{IC}}(\mathbf{x})] = n_0 [1 + \delta(\mathbf{x})]. \quad (44)$$

Due to the finite size of the geometry, \bar{n} fluctuates from one realisation to another such that $\bar{n} = n_0(1 + \epsilon)$, where

$$\epsilon := \frac{1}{V} \int_V \delta(\mathbf{x}) d^3\mathbf{x}. \quad (45)$$

As a result, one can relate the Gaussian field $\nu(\mathbf{x})$ to the corresponding Gaussian field $\nu_{\text{IC}}(\mathbf{x})$ affected by the IC directly with

$$\nu_{\text{IC}}(\mathbf{x}) = \frac{\nu(\mathbf{x}) - \epsilon}{1 + \epsilon}. \quad (46)$$

We can thus efficiently estimate the skewness of the power spectrum estimated in the Gaussian realisations with and without the presence of the IC. We found that without the IC, one can find an empirical model for the skewness of a Gaussian field in the cone geometries. Indeed, by defining an effective fundamental mode for a given geometry of volume V as $k_{\text{F,eff}} := 2\pi/V^{1/3}$, one can define an effective number of independent modes in a given k -shell as $N_{\text{eff}} := 2\pi k^2/k_{\text{F,eff}}^2$. Empirically we found that $S_{3,G}^{(0)} \simeq 2.2/\sqrt{N_{\text{eff}}}$ provides a good fit for both volumes we consider.

In Fig. 11 we show the skewness estimated from the COVMOS realisations for the two cones and the periodic box and compare them to the skewness estimated from the Gaussian realisations described above. We can see that the presence of a mask increases the level of skewness at all scales. The smaller is the mask the larger is the increase of the skewness level. This is expected because the application of the mask on the density field correlates Fourier modes. Indeed, this is in a way similar to increasing the size of the k -shell (see Sect. 4.2.2).

Our empirical model, which reproduces the skewness obtained from a Gaussian field without IC is shown in the top panel of Fig. 11 with dark solid lines. Comparing this empirical model, which does not include the effect of IC, to the skewness estimated from the Gaussian field with IC (dark solid lines against dotted lines in the top panel of Fig. 11), we observe that on small scales the IC itself is introducing an excess of skewness in the distribution of the power spectrum. We notice that its amount depends on the volume used to estimate the mean number density and that it is roughly constant at large k .

As a result, to quantify the excess skewness with respect to the Gaussian case, one needs to subtract the Gaussian skewness with IC to the estimated skewness in the COVMOS catalogues. This is shown in the lower panel of Fig. 11, where it appears clearly that when the considered volume is small ($0.06 [\text{Gpc}/h]^3$) it is important to take into account the skewness induced by the IC while when the volume is larger ($0.3 [\text{Gpc}/h]^3$) the absolute contribution from the IC to the total excess skewness becomes subdominant. As a reference, the smallest redshift bin for the *Euclid* spectroscopic sample of the first data release (DR1) is expected to be around $1.5 (\text{Gpc}/h)^3$, thus the IC should not increase the level of skewness in the *Euclid* two-point summary statistics in Fourier space.

5. Statistical distribution of individual Fourier modes within a shell

Up to now we studied the distribution of the shell-averaged estimated power spectrum and found that the excess of skewness with respect to a Gaussian distribution originates from the shell-averaged higher-order statistics, such as the trispectrum and pentaspectrum. We now want to understand the origin of this non-Gaussian distribution and whether it is generated by an intrinsic non-Gaussian distribution of each individual Fourier mode. In this section, we are thus interested in the distribution of square modulus of the density contrast $X_i := |\delta_{k_i}|^2$ evaluated at each individual wave mode within a shell k .

As shown in Appendix A, when δ_k is a Gaussian field then X_i is expected to follow an exponential distribution. We thus,

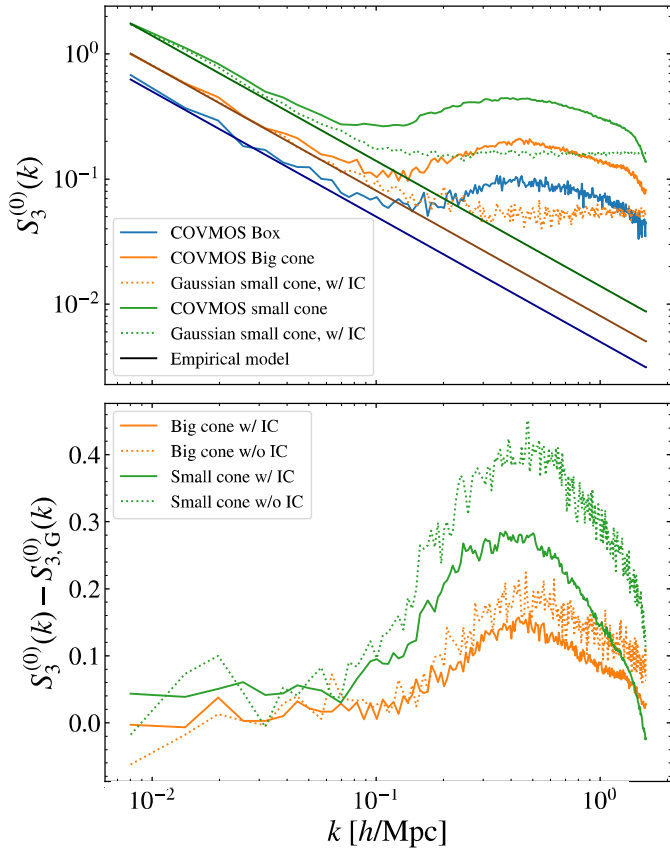


Fig. 11. Impact of the survey mask. *Top*: Estimated skewness in the three geometries from the COVMOS simulations (light coloured plain lines) and the Gaussian field realisations, where the contribution from IC is not subtracted (dotted lines). We also show our empirical model for the Gaussian field without the IC contribution (dark coloured solid lines). *Bottom*: Difference between the skewness estimated in the COVMOS simulations and the Gaussian realisations with and without IC (light coloured plain and dotted lines, respectively).

want to understand whether the skewness in the shell averaging process is generated by an intrinsic departure from the expected exponential distribution of each individual X_i or from the correlations between them. To achieve this we focus on the distribution of a sub-sample of modes within four Fourier shells $k = 0.2 h/\text{Mpc}$, $k = 0.3 h/\text{Mpc}$, $k = 0.4 h/\text{Mpc}$, and $k = 0.5 h/\text{Mpc}$. In each k -shell, we randomly select 1000 square modulus X_i , for which we can compute the one-point cumulant moments up to an order of five. We can compare these cumulant moments to the predicted ones assuming an exponential distribution (i.e. in the case of a Gaussian field). The latter can be expressed as

$$\langle X_i^n \rangle_c = (n-1)! \langle X_i \rangle^n, \quad (47)$$

as shown in Appendix A.

In Fig. 12 we display the relative deviation between the estimated cumulant moments of X_i with respect to their prediction of Eq. (47). Since for each shell we have 1000 square modulus X_i we only show the mean relative deviation (computed over the 1000 modes) and the corresponding dispersion over the considered shell. We can see a very good agreement between the estimation and its prediction in the case of an exponential distribution.

Despite the fact that the first five cumulant moments are well reproduced when fixing the first moment to the expectation value

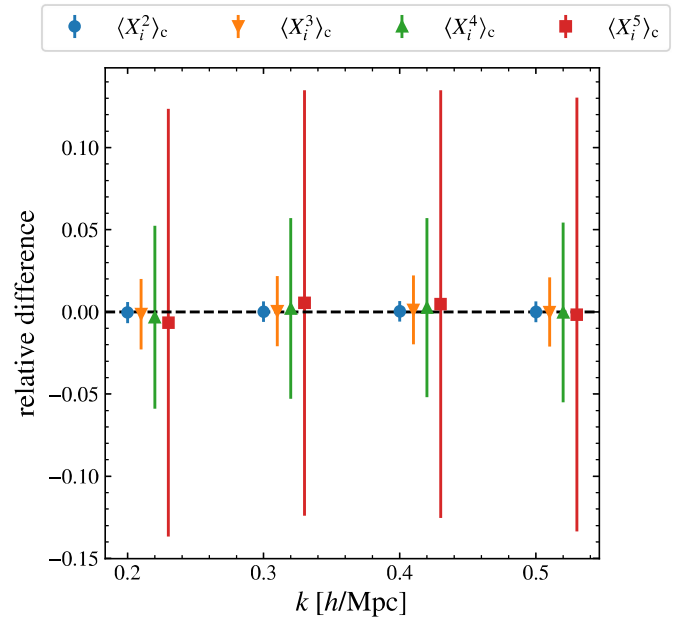


Fig. 12. Relative difference between the cumulants (up to an order of five) of each individual mode and their predictions in the case of a Gaussian field (see Eq. (47)). The error bars show the dispersion obtained for the 1000 considered wave modes.

evaluated over the 100 000 realisations, one could argue that this does not prove that the full distribution is well represented by an exponential distribution. Thus, we conducted a Kolmogorov–Smirnov (KS) test to compare each distribution of the 4×1000 wave modes with an exponential distribution. For each of the 4×1000 wave modes, we generate a set 100 000 values extracted from an exponential distribution with the estimated mean value and we perform the KS test to compare the true distribution and the exponential one. It turns out that the lowest probability assessing the inconsistency between the observed and expected distribution over all the tested modes is 0.73. It means that for none of the 4000 modes at our disposal we can reject the hypothesis that the distribution is exponential without having a very high probability of drawing a wrong conclusion. Thus, it appears that the square modulus at each measured wave mode is following with a great accuracy an exponential distribution.

In conclusion, this means that despite the fact that the density field is non-Gaussian, the distribution of the square modulus of the Fourier space density field is following the distribution which would be expected if the field were Gaussian. This is a clear indication that the main source of the excess skewness of the power spectrum distribution is due to the correlation between wave modes.

6. Conclusions

Given the level of precision of the cosmological constraints the *Euclid* mission aims at providing, the reliability of common assumptions that are usually made on the likelihood function must be checked to the same level of precision. One crucial assumption that was studied in detail in this paper is to consider that two-point statistics follow a Gaussian distribution.

We could test this assumption and compare it to our theoretical expectations with great precision thanks to the COVMOS method, which has been proven to be suitable to study

the covariance matrices of two-point statistics (Baratta et al. 2020, 2023; Gouyou Beauchamps et al. 2025) and the N -point distribution of two-point statistics in the present work (see the comparison with the DEUS–PUR simulations in Appendix B).

We first derived the analytical expressions of the second- and third-order cumulants of the power spectrum distribution in terms of the N -point functions of the density field. This allowed us to directly put in evidence of the relation between the non-Gaussianity of the density field and the non-Gaussianity of the power spectrum distribution and isolate the competing terms contributing to this distribution (see Eqs. (11) and (18)). In order to understand which terms were dominating depending on the scales, we derived the analytical expressions of the skewness and kurtosis of the multipoles of the power spectrum in both real and redshift space assuming a Gaussian field (see Eqs. (26)–(32) and Appendix A).

To compare these approximate predictions to their equivalent estimations for a more realistic non-Gaussian field, we generated several simulation samples, each containing 100 000 COVMOS realisations of the matter density field, while varying the redshift and the cosmological model both in real and redshift space. This allowed us to estimate the skewness and kurtosis of the distribution of the power spectrum monopole in real space and confirm, as was already observed in Takahashi et al. (2009) and Blot et al. (2015), the presence of an excess of skewness on non-linear scales ($k > 0.1 h/\text{Mpc}$ at $z = 0$). In addition, for the first time, we also estimated the skewness and kurtosis of the power spectrum multipoles in redshift space. We observed that the presence of RSD transfers the excess skewness of the monopole in real space to the multipoles but significantly reduces its overall amplitude (by roughly a factor of ten).

Based on our predictions and the estimation of the skewness in the COVMOS samples, we could demonstrate that the bispectrum contribution to the skewness (see Eq. (37)) is negligible because only a small fraction of the triangular configurations contribute to $\langle X^3 \rangle_c$. This led us to understand that on intermediate scales, the excess skewness is dominated by the trispectrum contribution, while on small scales it is dominated by the pentaspectrum of the matter density field.

We also extended this study to other cosmological models, including massive neutrinos and a time-varying dark energy equation of state. We have shown that varying the cosmological model does not have a significant impact on the skewness.

We then varied different settings in the estimation of the power spectrum to gauge their effects on its skewness. We observed that the aliasing due to the use of fast Fourier transforms has no impact on the skewness. However, varying the size of the k -shells of the power spectrum has a significant impact on the skewness, as increasing their size produces an increase in the excess skewness of up to 20%.

We also explored how the skewness was impacted by the number density of objects, which is a feature that is specific to a given survey. We saw that increasing the level of shot noise has the effect of reducing the skewness of the power spectrum on scales where the power spectrum is dominated by the shot noise. For a level of shot noise similar to the one expected for the *Euclid* spectroscopic sample, the skewness is greatly reduced (by around 50% for $k \sim 0.4 h/\text{Mpc}$, which are the most impacted scales).

However, all of this was done while assuming a periodic universe in a box, which is far from the reality of the data that will be obtained with *Euclid* (or any other survey). Indeed, we found that restricting the density field to a cone-shaped survey mask significantly increases the skewness of the power spectrum in

two different ways. Firstly, reducing the surveyed volume necessarily reduces the number of available Fourier modes, but it also correlates them, thus increasing the level of skewness at all scales. Secondly, the presence of the IC has a specific impact on the excess skewness on small scales, increasing it for smaller volumes.

Finally, in an attempt to understand the origin of the skewness of the power spectrum, we looked at the statistical distribution of each individual wave mode, $X = k_F^3 |\delta_k|^2$. We found, with a high level of accuracy, that they conserve an exponential distribution ($\theta^{-1} e^{-X/\theta}$, with parameter $\theta = k_F^3 \langle |\delta_k|^2 \rangle$) no matter their configurations. This led us to conclude that the extra skewness observed on small scales appears because of the correlation between wave modes within a k -shell.

To conclude, it is clear from this work that the distribution of two-point statistics is not Gaussian, and we understand the reasons for that. In particular, the shot noise, the scale binning, and the survey mask are factors that can significantly enhance its skewness. However, it is still unclear whether the assumption of a Gaussian likelihood, given non-Gaussian distributed data, would bias parameter inference for a dataset similar to what will be obtained with *Euclid*. This is explored in our companion paper GB-2025, where we show that these non-Gaussian features do not introduce a significant bias in cosmological parameter estimation for *Euclid*-like datasets.

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Appendix A: Skewness and kurtosis for the multipoles

In this appendix, we derive the expressions of the estimation of the skewness $S_3^{(\ell)}$ and kurtosis $S_4^{(\ell)}$ of the estimator of the power spectrum multipoles. We define the estimator X as a shell average at a fixed k modulus as

$$X := \frac{2\ell + 1}{N_k} \sum_{i=1}^{N_k} |\delta_{k_i}|^2 \mathcal{L}_\ell(\mu_i), \quad (\text{A.1})$$

where the sum runs over all modes contained in a given shell and μ_i is the cosine of the angle between the wave vector \mathbf{k} and the line-of-sight (along the z direction of the periodic box). Defining the variable $X_i := |\delta_{k_i}|^2$ and the coefficients $a_i^{(\ell)} := k_F^3(2\ell + 1)\mathcal{L}_\ell(\mu_i)/N_k$, one can express the estimator as a weighted average over the random variable X_i ,

$$X = \sum_{i=1}^{N_k} X_i a_i^{(\ell)}. \quad (\text{A.2})$$

Since the variable X_i is made of the sum of the square of the real and imaginary part of the Fourier modes δ_{k_i} , then it means that it follows an exponential distribution \mathcal{P} (Fisher et al. 1993) of parameter $\theta_i := P(\mathbf{k}_i)$, and thus

$$\mathcal{P}(X_i) = \frac{1}{\theta_i} e^{-X_i/\theta_i}. \quad (\text{A.3})$$

Its associated moment generating function $\mathcal{M}(t_i) := \langle e^{t_i X_i} \rangle$ is therefore given by

$$\mathcal{M}(t_i) = \frac{1}{1 - \theta_i t_i}, \quad (\text{A.4})$$

and the moment generating function of $Y_i := X_i a_i^{(\ell)}$ is given by $\mathcal{M}_Y(t_i) = \mathcal{M}(a_i^{(\ell)} t_i)$. Since we assume that each X_i variable is independent from each other, the same property holds for the Y_i variable, thus one can express the joint N_k -point moment generating function of the N_k variables as

$$\mathcal{M}_Y(t_1, \dots, t_{N_k}) = \prod_{i=1}^{N_k} \frac{1}{1 - \theta_i a_i^{(\ell)} t_i}. \quad (\text{A.5})$$

Thanks to the properties of the moment generating function, one can express the moment generating function of the sum of Y_i which is the variable X as $\mathcal{M}_X(t) = \mathcal{M}_Y(t, \dots, t)$ and its corresponding cumulant generating function $\mathcal{C}(t) := \ln\{\mathcal{M}_X(t)\}$ as

$$\mathcal{C}(t) = - \sum_{i=1}^{N_k} \ln(1 - \theta_i a_i^{(\ell)} t). \quad (\text{A.6})$$

One can take the n -th derivative of the above expression to obtain

$$\mathcal{C}^{(n)}(t) = (n-1)! \sum_{i=1}^{N_k} \frac{(\theta_i a_i^{(\ell)})^n}{(1 - \theta_i a_i^{(\ell)} t)^n}, \quad (\text{A.7})$$

from which it is straight forward to express the n -th order cumulant moments $\langle X^n \rangle_c = \mathcal{C}^{(n)}(t=0)$ as

$$\langle X^n \rangle_c = (n-1)! \sum_{i=1}^{N_k} (\theta_i a_i^{(\ell)})^n. \quad (\text{A.8})$$

Since the skewness S_3 and kurtosis S_4 are defined as

$$S_n := \frac{\langle X^n \rangle_c}{\langle X^2 \rangle_c^{n/2}}, \quad (\text{A.9})$$

we need to evaluate expression (A.8) for $n = 2, 3, 4$. Under some further hypothesis, one can obtain a closed analytical form. Assuming that the discrete sum over the shell can be approximated with its continuous limit, it follows that

$$\langle X^n \rangle_c \simeq \frac{(n-1)! (2\ell + 1)^n}{N_k^{n-1} 2} \int_{-1}^1 P^n(\mathbf{k}) \mathcal{L}_\ell^n(\mu) \mu. \quad (\text{A.10})$$

The above equation can be approximated when assuming that we are in real space, thus the power spectrum depends only on the modulus k of the wave vector \mathbf{k} :

$$\langle X^2 \rangle_c \simeq (2\ell + 1) \frac{P^2(k)}{N_k}, \quad (\text{A.11})$$

which is the usual expression for the variance of the power spectrum multipole estimator. For the third-order cumulant moment, we get

$$\langle X^3 \rangle_c \simeq (2\ell + 1)^3 2 \frac{P^3(k)}{N_k^2} \begin{pmatrix} \ell & \ell & \ell \\ 0 & 0 & 0 \end{pmatrix}^2, \quad (\text{A.12})$$

where we see the Wigner 3-j symbols, and finally for the fourth order,

$$\langle X^4 \rangle_c \simeq (2\ell + 1)^4 6 \frac{P^4(k)}{N_k^3} \sum_{n=0}^{2\ell} (2n+1) \begin{pmatrix} \ell & \ell & n \\ 0 & 0 & 0 \end{pmatrix}^4. \quad (\text{A.13})$$

This allowed us to express the skewness and kurtosis in the case of a Gaussian field as

$$S_{3,G}^{(\ell)} \simeq \frac{2}{\sqrt{N_k}} (2\ell + 1)^{3/2} \begin{pmatrix} \ell & \ell & \ell \\ 0 & 0 & 0 \end{pmatrix}^2 \quad (\text{A.14})$$

and

$$S_{4,G}^{(\ell)} \simeq (2\ell + 1)^2 \frac{6}{N_k} \sum_{n=0}^{2\ell} (2n+1) \begin{pmatrix} \ell & \ell & n \\ 0 & 0 & 0 \end{pmatrix}^4. \quad (\text{A.15})$$

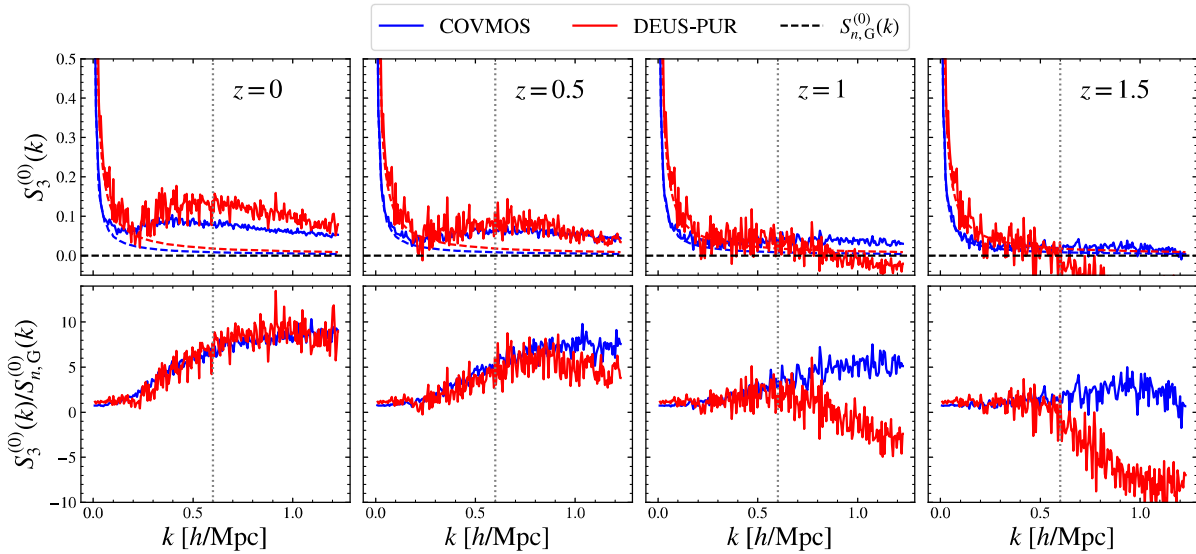


Fig. B.1. Comparison between COVMOS and DEUS-PUR. *Top:* Skewness of the distribution of the monopole of the real-space matter power spectrum (COVMOS and DEUS-PUR) in solid blue and red lines. The dashed line represents the corresponding predictions in the case of a Gaussian density field, computed using Eq. (26). *Bottom:* Ratio of the estimated skewness to the Gaussian field prediction, following the same colour code.

Appendix B: Validation of the COVMOS method against the DEUS-PUR N -body simulations

Here, we compare the results from the COVMOS realisations to those from the DEUS-PUR N -body simulation suite. The latter consists of 12 288 simulations with a volume of $656 \text{ (Mpc}/h)^3$ containing 256^3 tracer particles. All simulations have a Λ CDM cosmology characterised by the following set of parameter values $(h, \Omega_m h^2, \Omega_b h^2, n_s, \sigma_8) = (0.72, 0.1334, 0.02258, 0.963, 0.801)$. Using these simulations, Blot et al. (2015) have estimated the skewness of the monopole power spectrum estimator. We use their estimates to validate the skewness of the power spectrum we have obtained from the Λ CDM COVMOS realisations at $z = 0, 0.5, 1$ and 1.5 . It is important to remark that the DEUS-PUR estimates have been obtained with a binning in k that is set to half the fundamental frequency of the DEUS-PUR simulation box. Thus, for a given value of k , the number of independent modes may differ from the COVMOS case.

In Fig. B.1 we show the estimated skewness $S_3^{(0)}(k)$ at $z = 0, 0.5, 1$, and 1.5 from the COVMOS samples and from DEUS-PUR. We also show the Gaussian predictions for each case and the corresponding ratios, which exhibit the excess of non-Gaussianity due to the non-linearity of the matter density field. In plotting the excess non-Gaussianity of the COVMOS sample we have rescaled the ratio by a factor of $\sqrt{2}$, as we explain later in the text.

First of all, we can see that for each case the estimated skewness follows the Gaussian density field prediction on large scales. At $z = 0$ this occurs up to $k = 0.2 \text{ h}/\text{Mpc}$ and extends up to $k = 0.5 \text{ h}/\text{Mpc}$ at $z = 1.5$. At smaller scales, we can see an excess of non-Gaussianity due to the on-set of non-linearities that propagates from small to large scales for decreasing redshifts. It is worth noticing that in the case of the DEUS-PUR results, the estimated skewness drops at large k , with a greater effect at large redshifts. A similar trend can also be seen in the case of the COVMOS sample, though with a significant smaller amplitude. This is a numerical simulation resolution effect. In the case of the DEUS-PUR simulations, the effect is mit-

igated at smaller redshifts, where the resolution of the force calculation at small scales is improved by the adaptive-mesh refinement scheme. Indeed, if we conservatively consider only modes corresponding to scales larger than twice the coarse cell of the DEUS-PUR simulations, $k \lesssim k_{\text{Ny}}/2 = 0.6 \text{ h}/\text{Mpc}$, we are granted that the estimate skewness is free of such numerical systematic errors. The COVMOS samples, on the other hand, have been generated using power spectra estimated from the DEMNUni-Cov simulations that have greater spatial and mass resolution than the DEUS-PUR runs with a Nyquist frequency $k_{\text{Ny}} = 3 \text{ h}/\text{Mpc}$. That is why such an effect is much smaller than in the case of the DEUS-PUR estimates.

Yet, before comparing the excess non-Gaussianity of the COVMOS samples to that from DEUS-PUR, we should consider the effect of the different binning in k and different volume of the simulations. For this, we can use the approximation of the ratio of the excess skewness given by Eq. (42) since we are looking at non-linear scales ($k \gtrsim 0.2 \text{ h}/\text{Mpc}$ at $z = 0$). Indeed, given a binning Δk , we have that $N_k \simeq 2\pi k^2 \Delta k / k_{\text{P}}^3$, thus Eq. (42) reads as

$$\frac{S_3^{(0)}}{S_{3,G}^{(0)}} \simeq \frac{k \sqrt{2\pi\Delta k} \bar{Q}_0}{2 \bar{T}^{3/2}}. \quad (\text{B.1})$$

This shows that the simulation volume does not affect that excess skewness, only the binning in k . Given that the binning adopted for the COVMOS samples is twice as large as that of DEUS-PUR, the corresponding excess skewness differs by a factor of $\sqrt{2}$. In the bottom panels of Fig. B.1, we have shown the excess skewness of the COVMOS samples rescaled by a factor of $\sqrt{2}$ which agrees well with the estimates from DEUS-PUR over the range of non-linear scales which are unaffected by resolution effects. This validates the COVMOS analysis and demonstrates that such synthetic samples are able to reproduce the correct level of excess skewness expected from full N -body simulations.