

# External Validation of a Prognostic Model for Survival of Patients With Abdominal Aortic Aneurysms Treated by Endovascular Aneurysm Repair

Vaiva Dabravolskaitė<sup>a,b,†</sup>, Mometo M. Aweys<sup>c,‡</sup>, Maarit Venermo<sup>d</sup>, Harri Hakovirta<sup>b</sup>, Hozan Mufty<sup>e</sup>, Alexander Zimmermann<sup>d</sup>, Vladimir Makaloski<sup>a</sup>, Lorenz Meuli<sup>d,\*</sup>

<sup>a</sup> Department of vascular surgery, University of Bern, Inselspital, Bern, Switzerland

<sup>b</sup> Department of vascular surgery, University of Turku, Finland

<sup>c</sup> Department of Vascular Surgery, University of Helsinki and Helsinki University Hospital, Helsinki, Finland

<sup>d</sup> Department of vascular surgery, University of Zurich, Switzerland

<sup>e</sup> Department of Vascular Surgery, Leuven University Hospital, Belgium

## WHAT THIS PAPER ADDS

The validated prognostic model identifies a high risk subgroup of patients with asymptomatic abdominal aortic aneurysm (AAA) and a survival rate of only 16% at 10 years. The benefit of endovascular aneurysm repair in these patients must be questioned provided the AAA does not carry a relevant risk of rupture.

**Objective:** Current guidelines recommend diameter monitoring of small and asymptomatic abdominal aortic aneurysms (AAAs) due to the low risk of rupture. Elective AAA repair is recommended for diameters  $\geq 5.5$  cm in men and  $\geq 5.0$  cm in women. However, data supporting the efficacy of elective treatment for all patients above these thresholds are diverging. For a subgroup of patients, life expectancy might be very short, and elective AAA repair at the current threshold may not be justified. This study aimed to externally validate a predictive model for survival of patients with an asymptomatic AAA treated by endovascular aneurysm repair (EVAR).

**Methods:** This was a multicentre international retrospective observational cohort study. Data were collected from four European aortic centres treating patients between 2001 and 2021. The initial model included age, estimated glomerular filtration rate (eGFR), and chronic obstructive pulmonary disease (COPD) as independent predictors for survival. Model performance was measured by discrimination and calibration.

**Results:** The validation cohort included 1 500 patients with a median follow up of 65 months, during which 54.6% of the patients died. The external validation showed slightly decreased discrimination ability and signs of overfitting in model calibration. However, a high risk subgroup of patients with impaired survival rates was identified: octogenarians with eGFR  $< 60$  OR COPD, septuagenarians with eGFR  $< 30$ , and septuagenarians with eGFR  $< 60$  and COPD having survival rates of only 55.2% and 15.5% at five and 10 years, respectively.

**Conclusion:** EVAR is a valuable treatment option for AAA, especially for patients unsuitable for open repair. Nonetheless, not all these patients will benefit from EVAR, and an individualised treatment recommendation should include considerations on life expectancy. This study provides a risk stratification to identify patients who may not benefit from EVAR using the present diameter thresholds.

**Keywords:** Aortic aneurysm, Abdominal/surgery, Endovascular procedures/mortality, Predictive model, Risk factors, Proportional hazards models, Survival analysis

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## INTRODUCTION

Current treatment guidelines of the European Society for Vascular Surgery (ESVS) recommend diameter monitoring of

small and asymptomatic abdominal aortic aneurysms (AAAs) because the risk of rupture is very low.<sup>1</sup> Elective AAA repair is recommended for asymptomatic AAAs with diameters  $\geq 5.5$  cm in men and  $\geq 5.0$  cm in women (class I, level A for men; class IIb, level C for women).<sup>1–3</sup> However, robust data on the efficacy of elective treatment of all patients with an asymptomatic AAA with a diameter above this threshold are lacking, and the current level of evidence classification in men has been questioned.<sup>4</sup>

As the burden of comorbidities increases, elective AAA treatment becomes less effective or even futile in improving overall survival.<sup>4</sup> Thus, personalised decision making in

<sup>†</sup> D.V. and M.A.M. contributed equally to this study (shared first authorship).

\* Corresponding author. Department of Vascular Surgery, University Hospital Zurich (USZ), University of Zurich (UZH), Raemistrasse 100, CH-8091 Zurich, Switzerland.

E-mail address: [lorenz.meuli@usz.ch](mailto:lorenz.meuli@usz.ch) (Lorenz Meuli).

[@UszVascular](https://twitter.com/UszVascular)

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patients with asymptomatic AAA could avoid unnecessary AAA treatments and reduce the overall morbidity and costs associated with AAA. Nonetheless, this requires a reliable assessment of the impending risk of aneurysm rupture, the risks related to elective repair, and life expectancy. Only a complete picture of all competing risks will ultimately enable the benefit of endovascular aneurysm repair (EVAR) to be assessed individually.<sup>8</sup>

The aim of the current study was to evaluate the overall survival of patients treated by EVAR and thereby provide information on all cause mortality, the main adversary to the efficiency of EVAR in comorbid patients. The survival of patients with an asymptomatic AAA varies greatly after elective aneurysm repair.<sup>5–7</sup> Several patient characteristics and comorbidities have been associated with patient survival after elective AAA repair. Still, predicting survival after elective AAA surgery to support personalised decision making is not yet established due to the lack of robust and validated tools.

A single centre predictive model for survival after EVAR for AAA identified age, the estimated glomerular filtration rate (eGFR), and chronic obstructive pulmonary disease (COPD) as independent predictors for long term survival.<sup>9</sup> The previously published temporal validation demonstrated good discrimination ability for five year survival in four risk groups.<sup>7</sup> The five year survival probabilities were 89% in low risk patients, 83% in low to moderate risk patients, 68% in moderate to high risk patients, and only 40% in high risk patients. The current study aimed to externally validate this predictive model on an international multi-centre clinical cohort.

## MATERIALS AND METHODS

This retrospective observational cohort study includes all consecutive patients treated by standard EVAR for an asymptomatic AAA at four different European aortic referral centres: the university hospital of Zurich, Switzerland (2003 – 2020), the university hospitals of Turku and Helsinki, Finland (2010 – 2021 and 2002 – 2016, respectively), and the university hospital of Leuven, Belgium (2001 – 2019). Patients with complex EVAR, including fenestrated, branched, or parallel grafts, were excluded from this study. Further, all patients treated for symptomatic or ruptured aneurysms and other indications like penetrating aortic ulcers were excluded.

This study was conducted according to the Principles of the Declaration of Helsinki and reported in adherence to the TRIPOD statement (Transparent Reporting of a multi-variable prediction model for Individual Prognosis or Diagnosis).<sup>10</sup> The local ethics committees in Bern and Zurich Switzerland approved the study (BASEC-IDs: 2022-00489 and 2021-02311), whereas the local committees in Turku, Helsinki, and Leuven waived approval of the study due to its retrospective nature.

### Data collection and definitions

Patient characteristics were obtained from local records and aggregated at an individual patient level. Treatment indications were according to available ESVS guidelines from

2011 and 2018.<sup>1,11</sup> Prior to that, a threshold of 55 mm (50 mm for females), rapid progress ( $\geq 5$  mm in six months), or saccular anatomy was used as the indication criterion. Diameter measurements were extracted from recordings without consulting the available images.

The baseline characteristics of the validation cohort were summarised and compared with the original cohort. COPD was defined as any diagnosis of COPD at the time of operation or any forced expiratory volume  $< 80\%$  of the predicted capacity on pre-operative spirometry. eGFR was calculated using the modification of diet in renal disease study formula using the last pre-operative creatinine value within 30 days.

The primary outcome measure of this study was model performance measured by discrimination and calibration for overall survival in the validation cohort. For patients treated at the university hospitals of Leuven, Zurich, and Turku, survival information was obtained from local hospital databases. All patients without a documented date of death by the pre-defined study end date, 31 October 2022, were contacted during a cross sectional telephone survey between November 2022 and March 2023. For patients treated at the University of Helsinki, survival information was provided by the Statistics Finland Cause of Death registry. Completeness of follow up information was reported using the Follow up Index.<sup>12</sup> Survival information was trimmed at the study end date.

This overall dataset formed the validation cohort. The previously published patient cohort treated at the university hospital of Bern formed the original cohort and was used for comparison.

### Statistical analysis

**Predictor selection.** The variable selection process for the predictive score based on the original cohort and has been described in detail.<sup>9</sup> In summary, pre-selection of variables was conducted based on a literature review to avoid a complete data driven variable selection. Thereafter, a machine learning method (least absolute shrinkage and selection operator with 10 fold cross validation in a Cox model) was used for the variable selection for the predictive model. The predictive model identified age, eGFR, and COPD (see [Supplementary Table S1](#)).

**Predictive score.** The beta coefficients of the Cox model were used to create an easy to use risk score.<sup>7</sup> Age was grouped into quartiles and rounded to the next integer for practical reasons. eGFR was grouped into quartiles according to the KDIGO classification, but G4 and G5 were merged. COPD was available as a binary variable only and thus formed two groups.

The beta coefficients for each variable group were multiplied by 10 and rounded to the nearest integer to create the score. This resulted in the following scoring for age:  $< 70$  years = 0 points; 70 – 74.9 years = 9 points; 75 – 79.9 years = 10 points;  $\geq 80$  years = 17 points; for eGFR: KDIGO G1 = 0 points; G2 = 1 point; G3a = 3 points;

**Table 1.** Baseline characteristics of the aggregated validation cohort of patients with asymptomatic abdominal aortic aneurysm prior to endovascular aneurysm repair compared with original cohort

| Variable                          | Original cohort <i>n</i> = 552 | Validation cohort <i>n</i> = 1 500 | <i>p</i> value |
|-----------------------------------|--------------------------------|------------------------------------|----------------|
| Male sex                          | 503 (91.1)                     | 1 370 (91.3)                       | .88            |
| Age – <i>y</i>                    | 76.0 (69.4, 80.6)              | 75.2 (69.3, 80.0)                  | .44            |
| Arterial hypertension             | 458 (84.3)                     | 1 194 (80.6)                       | .055           |
| Missing                           | 9                              | 19                                 |                |
| Diabetes mellitus                 | 108 (19.6)                     | 252 (17.2)                         | .23            |
| Missing                           | 0                              | 39                                 |                |
| Dyslipidaemia                     | 467 (84.6)                     | 1 061 (72.6)                       | <.001          |
| Missing                           | 0                              | 38                                 |                |
| BMI – kg/m <sup>2</sup>           | 27.0 (24.0, 30.0)              | 26.1 (24.0, 29.2)                  | .004           |
| Missing                           | 29                             | 447                                |                |
| Smoking                           | 390 (73.3)                     | 855 (60.6)                         | <.001          |
| Missing                           | 20                             | 90                                 |                |
| COPD                              | 131 (23.7)                     | 470 (31.3)                         | <.001          |
| eGFR – mL/min/1.73 m <sup>2</sup> | 66.1 (51.0, 80.9)              | 77.8 (63.0, 87.5)                  | <.001          |
| Creatinine – mmol/L               | 91 (77, 111)                   | 93 (80, 112)                       | .036           |
| PAD, Fontaine class               |                                |                                    | <.001          |
| No PAD                            | 436 (79.0)                     | 496 (48.5)                         |                |
| Fontaine I                        | 49 (8.9)                       | 486 (47.6)                         |                |
| Fontaine II                       | 46 (8.3)                       | 38 (3.7)                           |                |
| Fontaine III                      | 21 (3.8)                       | 0 (0.0)                            |                |
| Fontaine IV                       | 0 (0.0)                        | 2 (0.2)                            |                |
| Missing                           | 0                              | 478                                |                |
| Coronary artery disease           | 308 (56.0)                     | 746 (50.1)                         | .019           |
| Missing                           | 2                              | 12                                 |                |
| Myocardial infarction             | 113 (20.6)                     | 353 (31.7)                         | <.001          |
| Missing                           | 4                              | 387                                |                |
| Aneurysm diameter – mm            | 58.0 (54.0, 62.5)              | 58.0 (55.0, 65.0)                  | .014           |
| Missing                           | 25                             | 7                                  |                |

Data are presented by median (quartiles 1, 3) or as *n* (%). Data were complete if not stated explicitly. Factor variables were compared by chi squared test, continuous variables by the Kruskal–Wallis rank test, respectively. BMI = body mass index; COPD = chronic obstructive pulmonary disease; eGFR = estimated glomerular filtration rate according to the Modification of Diet in Renal Disease Study (MDRD) in mL/min/1.73m<sup>2</sup>; PAD = peripheral arterial disease as clinical stage according to the Fontaine classification.

G3b = 6 points; G4/5 = 15 pts; for COPD: if present = 7 points.

The total score of the three variables was formed, and the cohort was divided into quartiles compiling the four risk groups: ≤ 8 points, low risk; 9 – 13 points, low to moderate risk; 15 – 18 points, moderate to high risk; ≥ 19 points, high risk. Of note, no combination resulted in a sum score of 14 points. The risk score is provided in [Supplementary Table S1](#) and is available online.<sup>7</sup>

**Model discrimination and calibration.** The discriminative ability of the model was tested on the validation cohort using Harrell's concordance statistics *C* with 95% confidence interval (95% CI) using DeLong's method.<sup>13</sup> Calibration was visually inspected and quantified using the jackknife pseudo value method. Discrimination and calibration of the predictive score were tested separately at five and 10 years. The same analysis was performed for the predictive model using age and eGFR as continuous rather than categorical variables. Further, the observed survival for each of the risk groups was compared between the original cohort and the validation cohort using the likelihood ratio test.

**Further descriptive analysis.** Continuous variables (i.e., age, eGFR, creatinine, body mass index, AAA diameter) were

visually inspected for normality and summarised using the median and quartiles (Q1, Q3) since they were skewed. Factor variables were compared by chi squared test and continuous variables by the Kruskal–Wallis rank test. A Cox proportional hazard model was calculated for the validation cohort, including all available variables previously identified to be associated with survival to allow the inclusion of these data in future studies. For this analysis, multiple imputations were performed for missing comorbidities using the Mice package. Predictive mean matching was used for continuous variables, multinomial logistic regressions were used to impute factor variables. The number of imputed datasets was *m* = 25. The proportional hazards assumption was tested and verified using scaled Schoenfeld residuals for each Cox model.

All statistical analyses were performed using R Studio version 4.2.3 on MacOS version 12.5.1.

## RESULTS

The four study sites treated a total of 1 616 consecutive patients with asymptomatic AAA using EVAR during their respective inclusion periods. Information on COPD was unavailable in 87 patients and eGFR was missing in 29 patients, resulting in a final validation cohort of 1 500

**Table 2.** Survival of the aggregated validation cohort of patients with asymptomatic abdominal aortic aneurysm prior to endovascular aneurysm repair compared with original cohort

| Variable                              | Original <i>n</i> = 552 | Validation <i>n</i> = 1 500 | <i>p</i> value |
|---------------------------------------|-------------------------|-----------------------------|----------------|
| Follow up – mo (Q1, Q3)               | 56 (23, 77)             | 65 (37, 101)                | <.001          |
| Follow up index                       | 0.95                    | 0.97                        | <.001          |
| Number of deaths                      | 197 (35.7)              | 819 (54.6)                  | n.a.           |
| <i>Overall survival</i>               |                         |                             |                |
| 30 days                               | 98.7 (97.8–99.7)        | 98.9 (98.3–99.4)            | .70            |
| 5 years                               | 70.8 (66.7–75.2)        | 70.8 (68.4–73.2)            | .70            |
| 10 years                              | 39.2 (32.3–47.6)        | 38.7 (35.7–42.0)            | .70            |
| <i>1 year survival by risk group</i>  |                         |                             |                |
| Low risk                              | 98.6 (96.6–100)         | 95.7 (93.8–97.7)            | .052           |
| Low to moderate                       | 98.6 (96.8–100)         | 94.6 (92.5–96.8)            | .33            |
| Moderate to high                      | 91.8 (87.0–96.8)        | 92.7 (90.1–95.3)            | .31            |
| High risk                             | 81.6 (75.3–88.6)        | 90.6 (87.2–94.1)            | .046           |
| <i>5 year survival by risk group</i>  |                         |                             |                |
| Low risk                              | 89.1 (83.7–94.9)        | 86.2 (82.7–89.8)            | .052           |
| Low to moderate                       | 83.7 (77.4–90.6)        | 74.0 (69.8–78.5)            | .33            |
| Moderate to high                      | 68.4 (59.5–78.6)        | 61.8 (56.9–67.2)            | .31            |
| High risk                             | 39.9 (31.7–50.2)        | 55.2 (49.3–61.8)            | .046           |
| <i>10 year survival by risk group</i> |                         |                             |                |
| Low risk                              | 74.8 (63.5–88.1)        | 61.2 (55.6–67.5)            | .052           |
| Low to moderate                       | 43.2 (30.1–62.0)        | 43.2 (37.6–49.6)            | .33            |
| Moderate to high                      | 23.4 (12.3–44.8)        | 24.6 (19.6–30.8)            | .31            |
| High risk                             | 9.6 (3.0–31.2)          | 15.5 (10.3–23.1)            | .046           |

Follow up time in months is summarized with median and quartiles (Q1, Q3); number of deaths with percentage (%); survival with 95% confidence intervals as given by the Kaplan–Meier estimators. n.a. = not available.

patients. Baseline characteristics of the aggregated validation cohort prior to EVAR are summarised in Table 1 and compared with the original cohort. Supplementary Table S2 summarises the baseline characteristics stratified by centre. The baseline characteristics of the validation cohort were comparable with the original cohort. However, significant differences were found for two variables of the predictive model. Patients in the validation cohort had a diagnosis of COPD more often (31.1% vs. 23.7%,  $p < .001$ ), and the eGFR was significantly higher in the validation cohort than in the original cohort (77.8 vs. 66.1 ml/min/1.73<sup>2</sup>,  $p < .001$ ).

The median overall follow up of the validation cohort was 65 (Q1, Q3: 37, 101) months with almost complete follow up information (Follow up Index 0.97). The overall survival times were identical among the two cohorts: 71% (95% CI 67–75) in the original and 71% (68–73) in the validation cohort; 39% (32–48) in the original and 39% (36–42) in the validation cohort ( $p = .70$ ), at five and 10 years, respectively. During follow up, 54.6% ( $n = 819$ ) of the patients in the validation cohort died, whereas 35.7% ( $n = 197$ ) of the patients in the original cohort died. Further details on follow up and survival at one, five, and 10 years are presented in Table 2.

### External validation of the predictive model

The external validation of the predictive model showed a slightly decreased discriminative ability compared with the previously reported model performance: Harrell's *C* 0.62 (95% CI 0.60–0.65) compared with 0.70 (0.66–0.75) on the original cohort. However, calibration was excellent: the

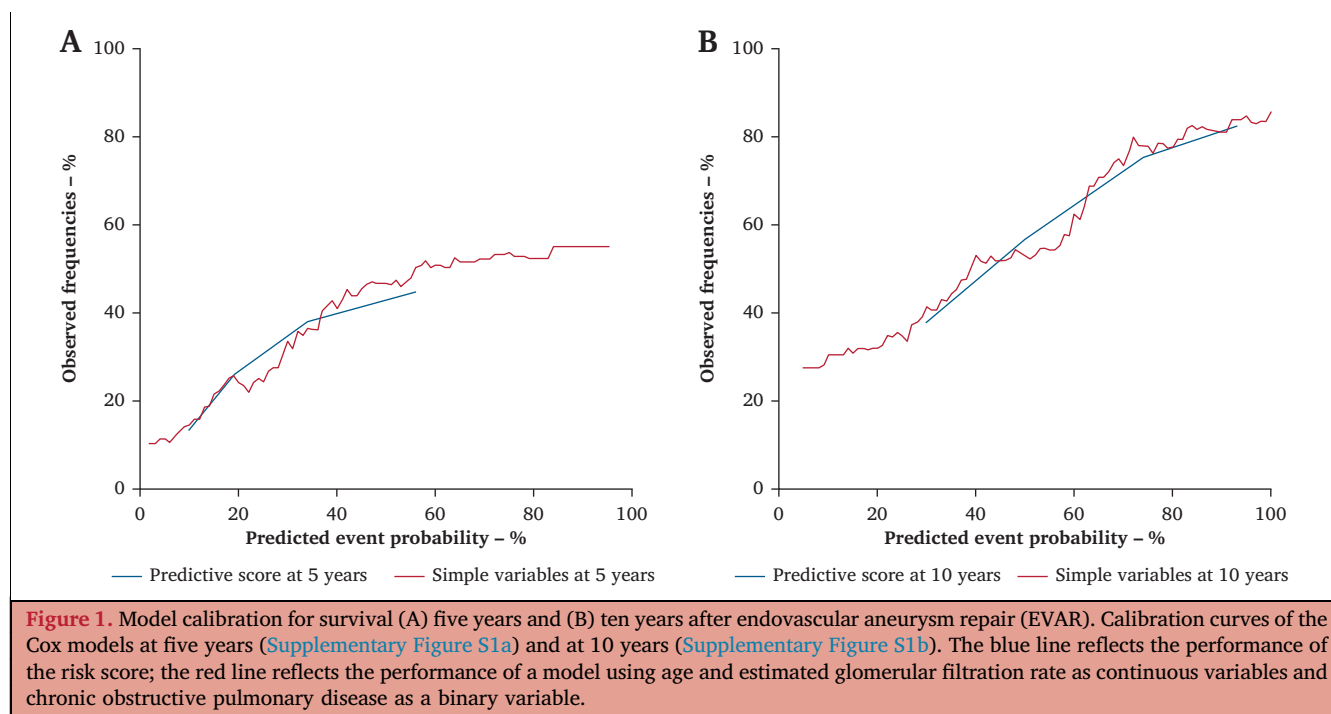
predicted and observed overall survival was 69.5% and 70.3% after five years and 37.0% and 38.3% after 10 years, respectively. The calibration curves at five and 10 years (Fig. 1) are slightly S shaped but very close to the perfect calibration reflected by the diagonal line.

Figure 2 and Table 2 show the survival at five and 10 years by risk group, comparing the original and external validation cohorts. There was no statistically significant difference in the observed survival between the original cohort and the validation cohort in the low risk group ( $p = .052$ ), the low to moderate risk group ( $p = .33$ ), and the moderate to high risk group ( $p = .31$ ). However, the observed survival was significantly better in the validation cohort compared with the original cohort for the high risk group ( $p = .046$ ). The five year survival in the original cohort for high risk patients was 40% (95% CI 32–50) and 55% (49–62) in the validation cohort. This demonstrates a slight but significant overfitting of the model for high risk patients with significantly better survival than predicted.

A multivariable Cox proportional hazard model for survival after EVAR, including all available variables in the dataset, is available in Supplementary Table S3.

### EVAR cohort over time

Figure 3 and Supplementary Figure S1 show changes in the risk scores and age groups of patients treated by EVAR in the validation cohort. A steady increase in the proportion of octogenarians can be seen from 2001, where only 8.8% were 80+ years old, to 2019, where 39.7% of the treated patients were octogenarians. Of note, there was a change in



this pattern with a decrease in the proportion of octogenarians and a decrease in the proportion of < 70 years old in the years 2020 and 2021 (Supplementary Figure S1).

The same observations were made for the burden of risk score. Patients were healthier, in terms of the risk score, with only 3.9% high risk patients in 2001, whereas 32% high risk patients in 2020. A steep drop in the proportion of high risk patients was documented for 2021, when only 11.1% were in the high risk group (Fig. 3).

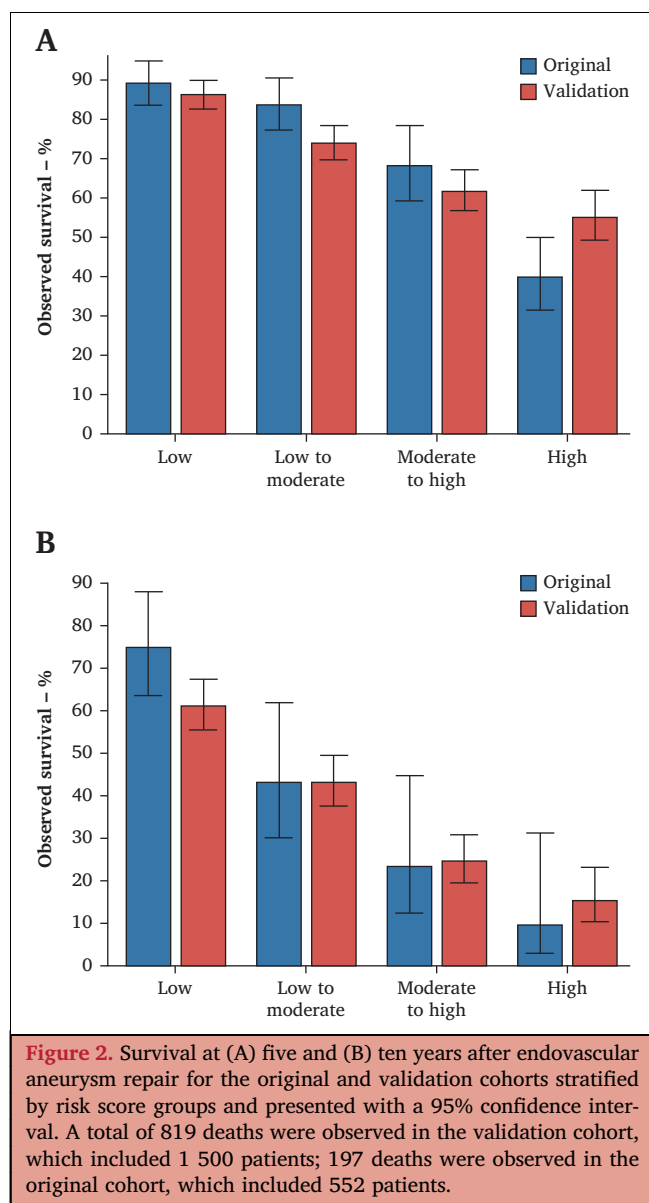
## DISCUSSION

This international, multicentre, external validation of a predictive model for the survival of patients with AAA treated by EVAR showed a modest reduction in the discriminatory ability, but excellent model calibration. The prognostic model confirmed a high risk subgroup of patients with a survival rate of only 55% at five and 16% at 10 years, respectively: octogenarians with eGFR < 60 or COPD, septuagenarians with eGFR < 30, and septuagenarians with both an eGFR < 60 and COPD. This contrasts with the excellent median life expectancy of the general population aged 80 in Switzerland, which in 2022 was 8.8 years for men and 10.4 years for women.<sup>14</sup> The benefit of EVAR in high risk patients must therefore be questioned provided the AAA does not carry a relevant risk of rupture.

Nonetheless, a complete picture of the risk–benefit balance of EVAR in this patient cohort should consider the following three points: (1) the median AAA diameter was 58 mm, about 1 cm smaller than in the patients who had participated in the EVAR 2 study.<sup>4</sup> The diameter of high risk patients was significantly larger ( $p < .001$ , Supplementary Figure S2). However, 50% of all patients in the high risk group had AAA diameters < 60 mm, and

approximately 25% had AAA diameters < 55 mm. The risk of AAA rupture might be lower than historical data suggest, but understanding the natural progression and rupture rates remains limited.<sup>15–17</sup> (2) The 30 day mortality rate after EVAR decreased but for these electively treated asymptomatic patients was still 1.1%.<sup>5</sup> (3) Like previous studies, this study shows that patients treated electively for an AAA have a relatively poor long term survival of only about 40% after 10 years.<sup>5,18–20</sup> Given this context, a substantial proportion of patients treated in this validation cohort may not have lived sufficiently long to realise the advantages of this preventive treatment. Strict adherence to the current diameter threshold in high risk patients or even expanding the treatment criteria for AAA patients beyond the binary threshold of 55 mm diameter would enhance the quality of patient care. Of note, an association between initial AAA diameter and survival after elective EVAR has been described previously.<sup>21</sup> However, AAA was eliminated in the variable selection process as the magnitude of this association was not strong enough.<sup>7,9</sup> Still, this association was confirmed in the multivariable analysis of this cohort, HR 1.01 per millimetre AAA diameter increase (95% CI 1.00–1.02,  $p < .001$ ); see Supplementary Table S3.

The decision regarding preventive treatment for an asymptomatic AAA in elderly and/or severely comorbid patients is challenging. The long term results of the EVAR 2 trial demonstrated no increase in overall life expectancy for the EVAR group vs. the non-treated group.<sup>6</sup> Of the originally included 404 patients in the EVAR 2 trial, only 17% (69/404) survived more than eight years, and these patients were younger, with higher body mass index, higher eGFR, and better forced expiratory volume in one second at the time of enrolment.<sup>6</sup> The 10 year survival rate of the high risk cohort in the current study is comparable to the overall



survival rate of patients from the EVAR 2 study. In contrast, improvements in peri-operative mortality rates have been achieved: a recent analysis from the American College of Surgeons National Surgical Quality Improvement Program of almost 25 000 patients undergoing EVAR between 2005 and 2013 showed a substantially lower 30 day mortality rate of 1.9% for high risk patients compared with the 7.3% reported by the EVAR 2 trial.<sup>6,22</sup> The peri-operative mortality rate in the high risk cohort was 2.2% and was comparable to these US data. Adkar *et al.* identified the presence of at least one criterion of impairment (respiratory, cardiac, or renal) or their combination as risk factors for 30 day death.<sup>22</sup> These risk factors for increased peri-operative mortality rate were comparable to the risk factors for long term survival identified in this study. Hence, individuals with elevated peri-operative risk have a reduced life expectancy, prompting the need to scrutinise the

advantages of elective AAA treatment based on the existing diameter threshold. Clinicians automatically and intuitively weigh different risks based on their experience and thus are likely to withhold EVAR in some patients. Diameters were already significantly larger in high risk patients compared with the other risk groups ( $p < .001$ ; [Supplementary Figure S2](#)). This study provides a risk stratification to support and improve such decisions in the future.

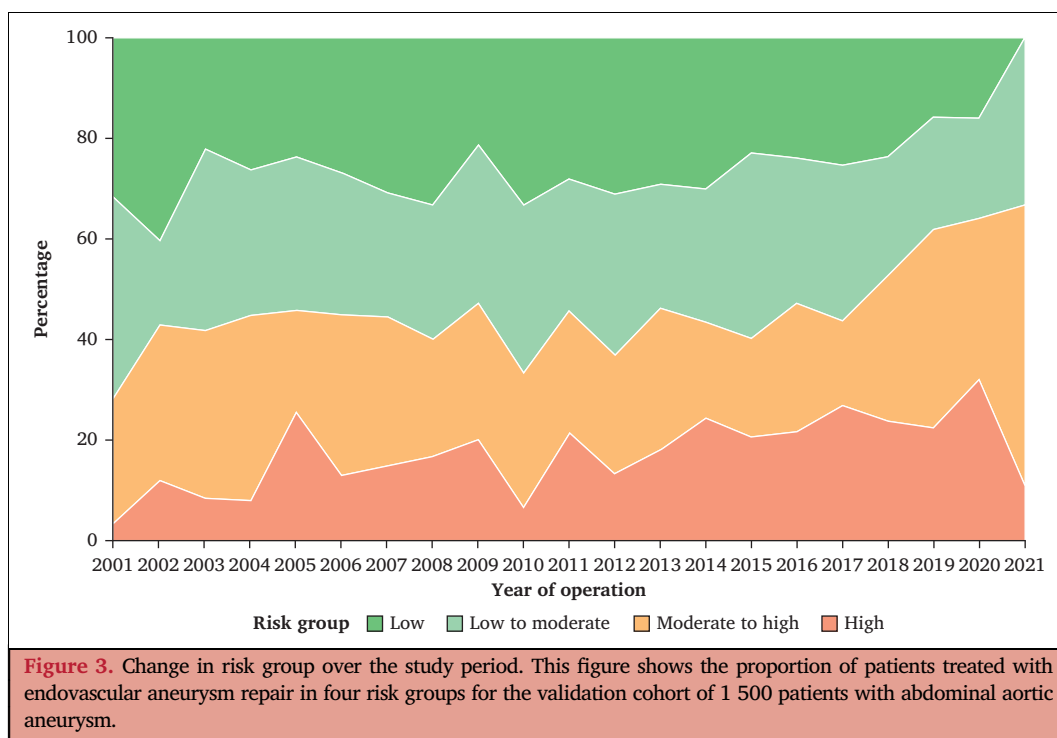
EVAR was initially invented as a less invasive alternative for patients unfit for open repair.<sup>23</sup> [Figures 2 and 3](#) depict the initial application of EVAR in relatively healthier patients, while its extensive use among elderly and more critically afflicted patients became predominant only during the last decade. This might be caused by a more restrictive use of EVAR in younger and healthier patients following long term results showing advantages for open repair.<sup>5,24</sup> A sharp decline in the proportion of high risk patients was documented for 2021, where only 11.1% were in the high risk group. This sharp decline falls within the COVID 19 pandemic and could be caused by reduced elective surgery capacity, especially for elderly and comorbid patients who are likely to require intermediate care or intensive care after treatment. It will be interesting to see if this trend continues in the years after COVID or if there is even a catch up effect.

The use of EVAR in low risk patients ( $< 70$  years with  $eGFR \geq 60$ , independent of COPD) who are expected to live longer is a separate topic of discussion. More than 60% of these patients will still be alive after 10 years, and thus at risk of late complications.<sup>5</sup> Primary open repair may still be the preferred treatment option for these patients. The coming years will show if there will be a trend towards an open first strategy for young low risk patients.

### Limitations

The model performance was validated and confirmed robust discrimination ability and excellent calibration to successfully identify a subset of high risk patients for impaired long term survival. The main limitations of this international multicentre external validation study are the retrospective extraction of routinely collected data which inherently carries a risk of bias: no routine pre-operative measurement of forced expiratory volume in one second was performed and COPD diagnosis was partly subjectively coded; furthermore, some heterogeneity and inconsistencies in AAA diameter measurement must be assumed over the two decades of the study period.

The calibration curve ([Supplementary Figure S1a,b](#)) and the survival plots show an S shaped model performance with better accuracy for the moderate risk groups. In contrast, the model slightly overestimates mortality in the high risk group and slightly underestimates mortality in the low risk group. This can be indicative of some degree of model overfitting or underfitting. Further validation of the model in cohorts with different case mix (i.e., higher or lower degrees of comorbidities) is needed to better understand calibration in the extremes of the calibration curve.



In general, the clinical applicability of any predictive model in daily routine needs to be carefully assessed. The idea behind risk stratification with predictive tools is to support clinical decisions rather than drive them.

### Conclusion

EVAR as a valuable treatment option remains undisputed, especially for patients unsuitable for open repair. Nonetheless, not all these patients will benefit from EVAR, and an individualised treatment recommendation should include considerations of life expectancy. This study provides a risk stratification to identify patients who may not benefit from EVAR under the present diameter threshold.

### CONFLICTS OF INTEREST STATEMENT AND FUNDING

None.

### APPENDIX A. SUPPLEMENTARY DATA

Supplementary data to this article can be found online at <https://doi.org/10.1016/j.ejvs.2023.11.018>.

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